

**DISCUSSION PAPERS
IN
ECONOMICS**

No. 2018/1 ISSN 1478-9396

**THE IMPACT OF UNCERTAINTY SHOCKS ON THE
VOLATILITY OF COMMODITY PRICES**

**DIMITRIOS BAKAS
ATHANASIOS TRIANTAFYLLOU**

JANUARY 2018

DISCUSSION PAPERS IN ECONOMICS

At Nottingham Business School, our Working Papers Series in Economics include a wide variety of approaches reflecting our research interests and academic expertise.

This paper is part of the new series, ***Discussion Papers in Economics***.

Earlier papers can be found at:

<https://www.ntu.ac.uk/research/research-at-ntu/academic-schools/research-at-nottingham-business-school/nbs-working-papers>

Please contact the Editor for enquiries concerning any of our Discussion Papers:

Marie Stack
Division of Economics
Nottingham Trent University
50 Shakespeare Street
Nottingham NG1 4FQ
UNITED KINGDOM

Email: Marie.Stack@ntu.ac.uk

Tel: + 44 (0)115 848 3880

The Impact of Uncertainty Shocks on the Volatility of Commodity Prices

Dimitrios Bakas ^{a,c} and Athanasios Triantafyllou ^{b†}

^a*Nottingham Business School, Nottingham Trent University, UK*

^b*Essex Business School, University of Essex, UK*

^c*Rimini Centre for Economic Analysis (RCEA), Canada*

Abstract

In this paper, we empirically examine the impact of uncertainty shocks on the volatility of commodity prices. Using alternative measures of economic uncertainty for the U.S. we estimate their effects on commodity price volatility by employing both VAR and OLS regression models. We find that the unobservable economic uncertainty measures of Jurado *et al.* (2015) have a significant and long-lasting positive impact on the volatility of commodity prices. Our results indicate that a positive shock in both macroeconomic and financial uncertainty leads to a persistent increase in the volatility of the broad commodity market index and of the individual commodity prices, with the macroeconomic effect being more significant. The impact is stronger in energy commodities compared to the agricultural and metals markets. In addition, our findings show that the measure of unpredictability of the macroeconomic environment has the most significant impact on the commodity price volatility when compared to the observable measures of economic uncertainty that have a rather small and transitory effect. Finally, we show that uncertainty in the macroeconomy is significantly reduced after the occurrence of large commodity market volatility episodes.

Keywords: *Economic Uncertainty, Commodity Prices, Volatility*

JEL Classification: *C22, E32, G13, O13, Q02*

† **Dimitrios Bakas:** Nottingham Business School, Nottingham Trent University, Burton Street, Nottingham, NG1 4BU, United Kingdom, *E-mail:* dimitrios.bakas@ntu.ac.uk.
Athanasios Triantafyllou: Essex Business School, University of Essex, Wivenhoe Park, Colchester, CO4 3SQ, United Kingdom. *E-mail:* a.triantafyllou@essex.ac.uk.

1. Introduction

Uncertainty shocks have a significant negative impact on the macroeconomy (Bloom, 2009; Colombo, 2013; Jurado *et al.*, 2015; Baker *et al.*, 2016; Caldara *et al.*, 2016; Henzel and Rengel, 2017; Meinen and Roeh, 2017). According to these empirical studies, a rise in economic uncertainty (as measured by several alternative proxies proposed in the literature) has a negative effect on aggregate investment, industrial production and the employment rate. Moreover, many recent empirical studies indicate that Economic Policy Uncertainty (EPU henceforth) shocks, in the form suggested by Baker *et al.* (2016), result to an increase in stock-market turbulence.¹ While many studies verify this negative impact of uncertainty shocks on the macroeconomy and on the equity markets, there is only limited empirical evidence in the literature about the impact of uncertainty shocks on the volatility of commodity markets (Watugala, 2015; Joets *et al.*, 2016; Van Robays, 2016).

In this paper, we attempt to extend the literature by examining the effects of economic uncertainty shocks on commodity market volatility.² Using various alternative proxies of economic uncertainty for the U.S. and a realized volatility measure for the broad commodity market index, as well as for a panel of 14 individual energy, agricultural and metal commodities, we estimate the impact of uncertainty on the commodity price volatility by means of both VAR and OLS regression analyses. All previous works (Watugala, 2015; Joets *et al.*, 2016; Van Robays, 2016) concentrate only on macroeconomic uncertainty without discriminating between

¹ Kang and Ratti (2014) find that economic policy uncertainty is interrelated to oil price shocks and has a negative impact in stock-market prices. Arouri *et al.* (2016) in their empirical analysis find that policy uncertainty shocks have a negative impact in stock-market returns, with the effect being more persistent in periods of extreme stock-market volatility. In further empirical support of these arguments, Antonakakis *et al.* (2013) show that the negative conditional correlation between economic policy uncertainty and stock-market returns and the respective positive conditional correlation between economic policy uncertainty and implied volatility (VIX index) is time varying and it is significantly affected by oil demand shocks and by the U.S. economic recessions. In addition, Liu and Zhang (2015) show that an increase in EPU has a positive impact on stock-market volatility and that EPU is a robust volatility forecasting factor for the stock-market, while Hattori *et al.* (2016) show that unconventional monetary policy announcements lead to a significant decrease in the option-implied perception of tail risks in the stock-market. Pastor and Veronesi (2012) find that the uncertainty about U.S. government policy decreases stock-market returns while at the same time increases volatility and risk premia in the stock-market. Lastly, Kelly *et al.* (2016) show that increasing political uncertainty significantly increases uncertainty in equity option markets, while Pastor and Veronesi (2013) show that political uncertainty increases volatility and makes stocks more correlated, especially during weak market conditions.

² With the term ‘economic uncertainty’ we refer to both macroeconomic and financial uncertainty. Later on in the paper we describe analytically the definitions of both macroeconomic and financial uncertainty shocks.

observable and latent uncertainty shocks, and do not provide any evidence about the impact of financial uncertainty. In this way, our paper is the first that explores a unified and more complete investigation of the impact of (observable and unobservable) macroeconomic and financial uncertainty shocks on commodity price volatility. Our primary motivation stems from the fact that the dynamics of commodity markets are driven by aggregate supply and demand (for commodities) shocks (Pindyck and Rotemberg, 1990; Kilian, 2009). Thus, when the timing or the magnitude of changes in macroeconomic conditions (e.g., industrial production) becomes more uncertain, as a result we expect turbulence in commodity markets to rise. Our secondary motivation stems from the relevant literature in commodity markets, according to which the booms and bursts in these markets are closely linked to macroeconomic fundamentals, such as the short-term interest rate and inflation (Frankel and Hardouvelis, 1985; Gordon and Rowenhorst, 2006; Frankel, 2008; Frankel and Rose, 2010; Gilbert, 2010; Gospodinov and Ng, 2013; Gubler and Hertweck, 2013).

We can identify in the literature two structurally different approaches for the measurement of economic uncertainty: observable and unobservable (or latent) uncertainty measures. The observable measures of economic uncertainty are those that can be proxied by the time-series variation of observable economic indicators, such as stock-market volatility (VXO) used in Bloom (2009) or uncertainty about future economic policy, which is based on economic news released in newspaper articles (EPU) (see Baker *et al.*, 2016, for more details on this approach). The unobservable economic uncertainty measures are based on the empirical method of Jurado *et al.* (2015) (and are referred hereafter as JLN measures). According to this approach, economic uncertainty cannot be measured by observed fluctuations in various economic indicators because these indicators may fluctuate for several reasons which are not at all related to uncertainty.³ Jurado *et al.* (2015) define and measure economic uncertainty as the volatility of the unforecastable component of a

³ Jurado *et al.* (2015) argue that stock-market volatility “*can change over time even if there is no change in uncertainty about economic fundamentals, if leverage changes, or if movements in risk aversion or sentiment are important drivers of economic fluctuations*”. This develops the limited earlier work analyzing these factors. For example, Abraham and Katz (1986) found that the time-variation in employment could be driven by heterogeneity of cyclicalities in the business activity of firms across the business cycle. Bekaert *et al.* (2013) show that aggregate economic uncertainty in the stock-market (as proxied by the VIX measure) is driven by changes in risk aversion and not by changes in economic uncertainty.

large group of important economic (macroeconomic and financial) indicators. In this paper, we use various alternative proxies for economic uncertainty in order to examine which type of uncertainty shock matters most for commodity investors. Our results reveal that a rising degree of unpredictability over the future state of the macroeconomy as well as of the financial sector (i.e., an increase in the unobservable JLN measures of Jurado *et al.*, 2015) is the most significant common factor of the contemporaneous rise in the volatility of commodity prices. The economic interpretation of this finding is that rising uncertainty about macroeconomic conditions is translated into rising uncertainty about future aggregate demand and supply, and since commodity prices are mainly driven by aggregate demand and supply conditions, their volatility increases due to these highly uncertain conditions in the macroeconomy. More specifically, our results show that the unobservable (latent) uncertainty JLN measures of Jurado *et al.* (2015) have a more significant and long-lasting impact on commodity market volatility compared to the observable economic uncertainty measures, such as the EPU index of Baker *et al.* (2016) and the VXO stock-market index. Therefore, what matters most for commodity investors, is not the macroeconomic and stock-market fluctuations per se, but the degree of unpredictability of these types of fluctuations. Commodity markets, according to our findings, are relatively immune to sudden ups and downs in the stock-market and to uncertainty about future economic policy. What is important for investors in commodity markets is their ability to foresee and anticipate the sudden swings and turbulence in the financial sphere and in the macroeconomy. As long as they achieve this, commodity markets become less volatile and less correlated with macroeconomic fluctuations.

Our econometric analysis reveals that in highly unpredictable times, commodity market volatility rises. This result may shed some light and provide a pure macroeconomic explanation of the rapid rise in the volatility of commodity prices over the 2006-2008 period. The analysis shows that the highly unpredictable macroeconomic environment (and not the highly volatile environment) is the key determinant of the rising volatility in the commodity markets. Our findings reveal that, the more that economic agents are able to forecast future macroeconomic fluctuations, the less volatile commodity markets will be. The policy implication behind these findings is that more transparent macroeconomic policies (e.g., a less discretionary monetary policy) can have a dual effect: the removal of the ‘fog’ and

uncertainty in the macroeconomy on the one hand, and the stability in commodity markets on the other.

In more detail, the VAR analysis shows that the unobservable economic uncertainty shocks have a more significant (in terms of magnitude) and long-lasting impact on the volatility of commodity prices compared to the observable uncertainty shocks. Our estimated Impulse Response Functions (IRFs) show that a 1% positive shock in the logarithm of the JLN uncertainty index increases the volatility in the commodity price index by 1.1% (in the case of a macroeconomic uncertainty shock) and by 0.6% (in the case of a financial uncertainty shock), with the responses of commodity market volatility remaining positive and statistically significant for almost 15 months after the initial uncertainty shocks. On the other hand, the impact of the EPU shocks on commodity market volatility has a much smaller and rather transitory effect on the volatility of commodity prices. Our estimated IRFs indicate that commodity market variance increases by 0.03% (3 basis points) after applying a 1% EPU shock and the response vanishes 2 months after the initial uncertainty shock. Our results do not change when we use the alternative components of the EPU index, for example the EPU news uncertainty, the Monetary Policy uncertainty proxy and the Fiscal Policy uncertainty proxy. Furthermore, our OLS regression models show that the estimated coefficients of the JLN macroeconomic uncertainty are larger and more significantly positive when compared to the estimated coefficients of the observable economic uncertainty measures, like the EPU index and its components, as well as the realized variance of the S&P 500 index and the level of the VXO implied volatility index.

The OLS regressions indicate significant explanatory power of the JLN macroeconomic uncertainty measure on the volatility of the commodity price index, with the adjusted R^2 values reaching almost 40%, while the respective R^2 value when using the EPU index falls to 7.7%. Furthermore, and despite the fact that monetary policy shocks have a significant negative impact on commodity prices and that an expansionary monetary policy is associated with higher commodity prices (see Frankel and Hardouvelis, 1985; Gordon and Rowenhorst, 2006; Frankel, 2008; Frankel and Rose, 2010; Gilbert, 2010; Anzuini *et al.*, 2013; Frankel, 2013; Gubler and Hertweck, 2013; Hammoudeh *et al.*, 2015), we find that the uncertainty about the future path of monetary policy has a rather transitory and insignificant impact on the volatility of commodity prices. Moreover, when examining the reverse channel of

causality, we find that the JLN macroeconomic uncertainty is significantly reduced after the occurrence of commodity volatility shocks. The significant reduction of macroeconomic uncertainty after the realization of large commodity volatility episodes, shows that the volatility of commodity prices represents a large fraction of the uncertainty in the macroeconomic environment, and when the volatility shock takes place, the future (expected) state of the macroeconomy becomes less ‘foggy’ as a result.

In addition to measuring the responses of uncertainty shocks on the volatility of the broad commodity futures index, we examine empirically the impact of economic uncertainty on a panel of individual commodities. In this analysis we include the most important (in terms of liquidity of the underlying commodity futures market) commodities for the energy, metals and agricultural commodity classes. Our main results and conclusions remain unaltered when we examine the impact of uncertainty shocks on the monthly realized variance of individual commodity futures prices. More specifically, we find that the volatility of both the agricultural, energy and metals commodity prices increases significantly after an uncertainty shock. The instant, synchronous and significantly positive jump of the volatility of commodity prices in response to the JLN macroeconomic uncertainty shocks, shows that macroeconomic uncertainty is a common (latent) factor behind the time-varying volatility of energy, metals and agricultural commodity markets. Moreover, our empirical analysis shows that the volatility of the energy commodity markets has a more instant and significant response to uncertainty shocks when compared to the volatility responses of the metals and agricultural commodity futures markets. Our results are in line with the findings of the relevant literature according to which oil price and uncertainty shocks have a negative impact on the macroeconomy (Ferderer, 1996; Hamilton, 2003; Kilian, 2008; Elder and Serletis, 2010; Rahman and Serletis, 2011; Jo, 2014; Elder, 2017). In addition, we identify also a reverse channel of causality, according to which, uncertainty about future economic activity affects significantly the volatility in energy commodity prices. Our work contributes to the relevant literature since we show that there is a causal nexus between uncertainty in the macroeconomy and in the oil market. While the empirical studies in the relevant literature show that higher uncertainty in crude oil markets depresses economic activity, we additionally show that higher uncertainty in the macroeconomy creates more turbulence in the oil market.

Previous empirical work in the literature does not provide any support for a common macroeconomic factor driving the time-variation in commodity market volatility. The volatility dynamics in commodity markets are mainly attributed to commodity market fundamentals which are based in the ‘Theory of Storage’ (see Kaldor, 1939; Working, 1948, 1949; Brennan, 1958; Telser, 1958; Fama and French, 1987). According to this theory, commodity market turbulence occurs due to rising convenience yields and subsequent commodity market fears about a stock-out in inventory levels. All empirical studies exploring the common factors driving volatility in commodity markets, conclude that there is no common macro factor behind the time-varying volatility in commodity prices.⁴ Our results do not contradict these empirical findings, but provide a macroeconomic ‘explanation’ of this inverse relationship between convenience yields, inventory levels and volatility in commodity prices, as implied by the ‘Theory of Storage’. Rising macroeconomic uncertainty makes the path of future aggregate demand (and consequently of aggregate production) less predictable. Therefore, risk averse commodity producers will prefer to hold physical inventory when facing uncertain aggregate demand conditions.⁵ This rise in the convenience yield for holding physical inventory will result in a rapid rise in the volatility of the commodity prices, confirming the ‘Theory of Storage’ (Ng and Pirrong, 1994; Fong and See, 2001).

Finally, we perform the same OLS regression and VAR analysis over the post-2000 period, during which financialization of commodity markets has greatly increased. Thus, we examine the consequences of the financialization process, and explore, at the same time, whether the dynamic interactions between commodity market volatility and economic uncertainty have changed in magnitude over this period. Our results indicate that the impact of uncertainty shocks has increased

⁴ For example, Batten *et al.* (2010) show that there are no common macroeconomic factors influencing the dynamics of the monthly volatility series of metals prices. According to their findings the monthly volatility of gold prices is affected by changes in monetary factors, while the same is not true for silver. In further support of these empirical results, Hamoudeh and Yuan (2008) find that while the monetary and oil price shocks reduce the volatility of precious metals (gold and silver), they do not have the same effect for the volatility of copper prices. In addition, Pindyck and Rotemberg (1990) were among the first to identify the “excess co-movement” of commodity prices. They show that this excess co-movement is well in excess of anything that can be explained by common macroeconomic factors like inflation, exchange rates or changes in aggregate demand. Our empirical findings may provide an explanation for this “puzzling phenomenon” of Pindyck and Rotemberg (1990), since we empirically verify that macroeconomic uncertainty is a common macroeconomic factor which lies behind the time-variation in commodity market price volatility.

⁵ More recent empirical findings (Gospodinov and Ng, 2013) show that the convenience yields of commodity markets are significant indicators of future inflation. Specifically, they attribute the forecasting power of the convenience yields on their ability to indicate future economic conditions.

exponentially during the post-financialization era. This shows that the financialization process, in addition of the transformation of the commodity markets into a separate asset class, has increased the structural interconnections between macroeconomic and commodity markets uncertainty.

The rest of the paper is organized as follows. **Section 2** describes the data and outlines the empirical methodology. **Section 3** presents the empirical results, while **Section 4** provides various robustness checks. Finally, **Section 5** concludes.

2. Data and Empirical Methodology

2.1 Commodity Futures Data

We use the daily excess returns of the S&P GSCI indices on commodity futures prices. More specifically, we use the daily excess returns of the broad commodity futures market index as our proxy for the daily price of commodities. In addition, we obtain the individual daily time series of agricultural, energy and metals commodities of the S&P GSCI commodity futures indices. Our cross-section of agricultural commodities includes cocoa, corn, cotton, soybeans, sugar and wheat, while the cross-section of energy commodities includes crude oil, heating oil, petroleum and unleaded gasoline, and lastly, the cross-section of metals commodities includes gold, silver, copper and platinum. The commodity futures dataset covers the period from January 1988 until December 2016. All the S&P GSCI daily series of commodity futures prices are downloaded from Datastream.

For the estimation of the monthly realized variance we follow the empirical approach of Christensen and Prabhala (1998) and Wang *et al.* (2012) and we estimate the realized variance as the monthly variance of the daily returns of commodity futures as follows:⁶

$$RV_{t,T} = \frac{1}{T} \sum_{i=1}^T \left(\frac{F_{t+i} - F_{t+i-1}}{F_{t+i-1}} - \frac{\overline{F_{t+i} - F_{t+i-1}}}{F_{t+i-1}} \right)^2, \quad (1)$$

⁶ In the financial econometrics literature the realized variance (RV) is usually defined as the best discrete time estimator of the quadratic return variation (QV) which is equal to the sum of the quadratic realized returns (Carr and Wu, 2009) for a given time period. This kind of estimation is usually applied when dealing with high-frequency intraday data for which the sum of quadratic returns converges more efficiently to the integrated quadratic return variation process (Barndorff-Nielsen and Sheppard, 2002). In our case, where we deal with daily data, we follow the empirical approach of Wang *et al.* (2012) by estimating the realized variance as the monthly variance of daily commodity futures returns for the given month, as shown in Equation (1).

where F_t is the commodity futures price the trading day t and the time interval (t, T) is the number of trading days during each monthly period. $RV_{t,T}$ is the estimated realized variance for the each monthly period. Our monthly estimate of the annualized realized variance ($COMRV$) is the monthly variance of the daily returns of commodity prices (for each monthly period), multiplied by 252 (the number of trading days in each calendar year) in order to be annualized ($COMRV = RV_{t,T} * 252$).

2.2 Economic Uncertainty Data

The unobserved (latent) measures of macroeconomic uncertainty (MU) and Financial Uncertainty (FU) are based on the JLN approach.⁷ The JNL MU1, MU3 and MU12 uncertainty measures are the macroeconomic uncertainty series which represent the unobservable estimate of macroeconomic uncertainty when having 1, 3 and 12 month forecasting horizon respectively. The same holds for the JLN FU1, FU3 and FU12 financial uncertainty variables. In our main econometric analysis, we use the 3-month ahead macroeconomic uncertainty (MU3) and financial uncertainty (FU3) measures as the benchmark cases. Thus, our MU and FU time series correspond to the MU3 and FU3 uncertainty series of JLN.⁸

Our observable measure of economic uncertainty is based on the approach of Baker *et al.* (2016), according to which economic uncertainty is proxied by the uncertainty about economic policy which can be observed in economic news and newspaper articles. We include the monthly time series for the EPU index by Baker *et al.* (2016), and its components, containing the fiscal policy uncertainty (EPUFISC), monetary policy uncertainty (EPUMON), as well as the uncertainty measure on news about economic policy (EPUNEWS) and the uncertainty measure on news about financial regulation (FRU).⁹ We additionally include some widely accepted measures of economic uncertainty (see Bloom, 2009), like the monthly VXO index and the Realized Variance of the daily returns of the S&P 500 stock-market index (SP500RV). The daily series of the S&P 500 index (SP500) and the monthly VXO

⁷ The measures of Jurado *et al.* (2015) are downloaded from: <https://www.sydneyludvigson.com/data-and-appendixes>.

⁸ For robustness purposes, we provide additional results using the JNL MU1, MU12 and FU1, FU12 measures of economic uncertainty. The empirical findings using these measures of uncertainty remain unaltered.

⁹ The measures of Baker *et al.* (2016) are downloaded from the EPU website at: <http://www.policyuncertainty.com>.

data are downloaded from Datastream. All the economic uncertainty series have monthly frequency and cover the period from January 1985 until December 2016.¹⁰

2.3 Macroeconomic Data

We obtain monthly time series data for the U.S. Industrial Production Index and Employment in the Manufacturing Sector (MIPI and MEMP). We additionally use the U.S. effective exchange rate (EXCH). We lastly estimate the slope of the term structure (or term spread) as the difference between the 10-year constant U.S. government bond yield and the 3-month U.S. Treasury Bill rate (TERM), and we include also the logarithm of the oil price (OILP). All the monthly macroeconomic time-series variables used in our analysis are downloaded from Datastream and cover the period from January 1985 until December 2016.

3. Empirical Results

3.1 Descriptive Statistics and Unit Root Tests

Table 1 presents the descriptive statistics along with the respective unit-root tests of the commodity price volatility, the various uncertainty measures and the financial and macroeconomic control variables which are used in the empirical analysis.¹¹

[Insert Table 1 Here]

From **Table 1** we observe that all explanatory variables are stationary (both the ADF and the PP unit root tests reject the hypothesis of a unit root for all variables). The only exceptions are the log of the manufacturing Industrial Production index and the log of the U.S. Employment in the manufacturing sector for which we cannot reject the hypothesis of a unit root. For this reason, we continue using the first differences of these variables in our OLS regression models, which, according to

¹⁰ The Realized Variance (SP500RV) of the S&P 500 index has been estimated by applying the same methodology as in Equation (1).

¹¹ The time series variables of unobserved macroeconomic and financial uncertainty (namely the variables MU1, MU3, MU12, FU1, FU3, FU12) have been multiplied with 100 in order to be comparable with the observable economic uncertainty measures like the EPU level and its components. This transformation is essential in order to measure and compare the magnitude of the impact between observable (EPU) and unobservable (MU and FU) uncertainty shocks. Using this transformation, the estimated OLS coefficients as well as the estimated Impulse Response Functions (IRFs) based on the VAR models are of the same magnitude, and thus their impact can be directly comparable.

Table 1, are found to be stationary processes.¹² Furthermore, the means of the logarithmic uncertainty measures have nearly equal values (for example, the mean value of the log of the EPU index is 4.639 while the mean of the log of the MU index is 4.356), while on the other hand, the volatility (i.e., the standard deviation) of the observable economic uncertainty indices like the VXO and the EPU index is nearly ten times larger compared with the standard deviation of the unobservable JLN MU and FU indices. Moreover, **Figures 1** and **2** below show the contemporaneous time series movements of the Realized Variance (COMRV) of the commodity price index and the MU and FU uncertainty series respectively.

[Insert Figure 1 Here]

[Insert Figure 2 Here]

Figure 1 shows the contemporaneous movements of the JLN MU and the Realized Variance (COMRV) in commodity prices. We can observe that rapid rises in MU are being followed by jumps for the variance of the commodity price index and that the realizations of large shocks in the RV of commodity prices (e.g., the 2008-2009 volatility episode) are being followed by less uncertain (as indicated by a rapid reduction in the MU index) macroeconomic environment. **Figure 2** shows the respective contemporaneous movements of the JNL FU index and the variance of the commodity price index. The relationship between FU and commodity price RV is similar with that of MU and RV, but there are some increases in the FU index which are not being followed by analogous jumps for the RV of commodity price index.

3.2 The Impact of Uncertainty Shocks on Commodity Price Volatility

3.2.1 The Impact of Unobservable Uncertainty Shocks

In this section we present the results of our multivariate (6-factor) VAR model in which we include as endogenous variables the logarithm of the Manufacturing Industrial Production Index (MIPI), the logarithm of the Manufacturing U.S. employment (MEMP), the logarithm of the uncertainty index (log(Uncertainty) – MU, FU or EPU accordingly), the term spread (TERM) (the difference between the

¹² In the appendix we additionally employ the same OLS regression models using the log of the manufacturing Industrial Production and the Employment rate series, despite being $I(1)$ processes, for robustness.

10-year U.S. government bond yield and the 3-month U.S. T-bill rate), the logarithm of the monthly price S&P500 index (SP500RV) and the Realized Variance of the daily returns of Commodity Futures price index (COMRV)).¹³ The estimated 6-factor model is inspired by the multivariate VAR models of Bloom (2009) and Baker *et al.* (2016).¹⁴

The reduced form VAR model is given in Equation (2) below:

$$Y_t = A_0 + A_1 Y_{t-1} + \dots + A_k Y_{t-k} + \varepsilon_t, \quad (2)$$

where A_0 is a vector of constants, A_1 to A_k are matrices of coefficients and ε_t is the vector of disturbances which have serially uncorrelated disturbances, zero mean and a variance-covariance matrix $E(\varepsilon_t, \varepsilon_t') = \sigma_\varepsilon^2 I$. Y_t is the vector of endogenous variables. All variables are in monthly frequency and cover the period from January 1985 until December 2016. The ordering in our 6-factor VAR model is as follows:

$$Y_t = [MIPI_t, MEMP_t, MU_t, TERM_t, SP500_t, COMRV_t]. \quad (3)$$

Following the modeling approach of Bekaert *et al.* (2013), we choose to place macroeconomic variables first and the financial variables (term spread, stock-market, and commodity market) last in the VAR ordering selection due to more sluggish response of the former compared to the latter ones. We estimate a VAR model with 4 lags ($k=4$ in Equation (2)). The VAR(4) model is selected based on the Frechet and the Akaike optimal lag-length VAR criteria.¹⁵ **Table 2**, reports the Granger causality

¹³ With the term ‘uncertainty’ we denote all the alternative economic uncertainty indices we employ in order to measure the impact of the different indicators of economic uncertainty. Economic uncertainty refers both to macroeconomic and financial uncertainty. In our empirical analysis, in the main paper and the online appendix, we use six different indicators of economic uncertainty and five different indicators of financial uncertainty, thus, we estimate a total of ten multivariate VAR models.

¹⁴ Since we want to examine the impact of uncertainty shocks on the commodity price volatility, and since the commodity prices are directly linked to the manufacturing production process, we choose to include the Manufacturing Industrial Production and Employment (instead of the respective aggregate figures for U.S. Industrial Production and Employment which are being used in the VAR model of Bloom (2009) and Baker *et al.* (2016)). Another minor difference is that instead of the Federal Funds Rate (FFR) and the logarithm of the Consumer Price Index (log(CPI)), we use the term spread which includes the expectations about the future level of short-term interest rates and inflation. In addition, in our baseline VAR model we exclude the wages and the working hours.

¹⁵ We additionally run the VAR(3) model in order to compare the results with the VAR(3) model of Baker *et al.* (2016). When estimated a VAR model with 3 instead of 4 lags, the main results and conclusions remain unaltered. The estimated IRFs for the VAR(3) model are provided upon request.

tests between the alternative proxies of economic uncertainty and the volatility of the commodity price index. The tests are conducted using the baseline 6-factor VAR model given in Equation (3), in order to control for different macroeconomic and financial shocks like the industrial production and the interest rates (term spread) shocks. We estimate a total of nine VAR models by placing the 9 alternative economic uncertainty proxies as the third variable of the VAR ordering (in the place of MU as shown in Equation (3)).

[Insert Table 2 Here]

From **Table 2** we can observe that almost all proxies for economic uncertainty Granger cause the Realized Variance of the commodity price index (COMRV). More specifically, the Financial Uncertainty (FU), the Macroeconomic Uncertainty (MU) and the Economic Policy Uncertainty (EPU) (and its main components) Granger cause the volatility of the commodity price index. In addition, the causality tests reveal a bi-directional causal relationship between commodity market volatility and the JLN Macroeconomic Uncertainty (MU) measure. The changes in the RV of commodity price index cause changes in MU. On the other hand, as Panel C of **Table 2** indicates, while the volatility in the S&P 500 index and the financial regulation uncertainty (FRU) index have a causal effect in commodity market volatility, we fail to reject the hypothesis of no causality when conducting the test between the logarithm of the VXO index and the realized variance of the commodity market index (COMRV).

We continue the empirical analysis by measuring the impact of uncertainty shocks on the volatility of commodity prices. The impact of uncertainty shocks is quantified by estimating the Impulse Response Functions of the multivariate VAR model presented in Equation (3). More specifically, we base our analysis on the estimated IRFs between the logarithm of the various uncertainty indices and the realized variance of the commodity futures price index (COMRV). We firstly estimate the 6-factor VAR in which we use the unobserved proxy for macroeconomic uncertainty, i.e., the MU measure. The estimated Impulse Response Functions (IRFs)

between Macroeconomic Uncertainty (MU) and the Realized Variance of the Commodity Futures price index (COMRV) are given in **Figure 3**.¹⁶

[Insert Figure 3 Here]

The estimated IRFs in **Figure 3** show that a one percentage point (1%) shock in Macroeconomic Uncertainty (MU) raises the monthly variance of the commodity price index (COMRV) by almost 1.1% for the first 3 months after the initial shock with the effect being positive and statistically significant for almost 15 months after the initial shock of uncertainty. In other words, we find that an increase in the JLN unobservable MU measure has a tremendous and long-lasting impact on the volatility of commodity prices. According to our VAR analysis, we are the first that provide evidence on the existence of common macroeconomic uncertainty factors which drive the dynamics of the time-varying volatility in commodity prices.¹⁷ Our results are in sharp contrast with the findings of the relevant literature (for example, Batten *et al.*, 2010), according to which there are no macroeconomic factors who can jointly influence the volatility of the commodity price series. On the contrary, our analysis reveals that the MU factor is a significant determinant of time-variation in the broad commodity futures price index.

When examining the reverse channel of causality, we find that a positive shock in the realized variance (COMRV) of commodity prices reduces macroeconomic uncertainty (MU) in the short to medium run (2 months after the initial commodity volatility shock). The negative effect reaches its maximum 10 months after the initial commodity volatility shock and remains significant (i.e., statistically different from zero according to the bootstrapped 95% confidence interval) for about 25 months. The

¹⁶ In our online appendix we report the estimated orthogonalized Impulse Response Functions in which the shocks in the VAR model are orthogonalized using a Cholesky decomposition. According to Pesaran and Shin (1998), the generalized IRFs are invariant to the ordering of the variables in the VAR model, while the OIRFs are highly sensitive to the VAR ordering. For this reason, in our robustness section, we report the estimated OIRFs for different VAR orderings of the endogenous variables including in the VAR system. Following Pesaran and Shin (1998), we do not have to report the results of the estimated reduced form IRFs for different VAR orderings, since these IRFs are VAR ordering invariant. In addition, our results and basic conclusions remain unaltered when we estimate the OIRFs instead of the generalized ones. Koop *et al.* (1996) and Diebold and Yilmaz (2012) give further empirical support on these findings.

¹⁷ In the next paragraphs, we provide further robustness and empirical support to this finding. We show that the MU shocks have a significant and long-lasting impact on the volatility of individual commodity prices (e.g., minerals and agricultural products), and not only on the variance of the broad commodity price index.

more sluggish response of macroeconomic uncertainty to changes in the volatility of commodity prices is somewhat expected. The economic interpretation of this negative response has its roots in the construction of the unobservable JLN macroeconomic uncertainty index. This index has been estimated as the purely unforecastable component of macroeconomic fluctuations, thus, when a large commodity volatility episode is materialized, the uncertainty (or the degree of unpredictability) in the macroeconomy falls because of the realization of the highly uncertain shock in the commodity markets. Our VAR analysis indicates that when a large shock in commodity markets materializes, then a large fraction of the foggy and uncertain state of the future path of the macroeconomy disappears. Our empirical findings show, for the first time in the literature, that the increasing volatility in commodity markets has significant bi-directional linkages with the time-varying degree of unpredictability in macroeconomic fluctuations.

Moreover, we estimate the same 6-factor VAR model of Equation (1) with the Financial Uncertainty (FU) index instead of the Macroeconomic Uncertainty (MU) index. The Financial Uncertainty (FU) index (just like the MU index) is an unobservable uncertainty index which measures the degree of unpredictability in financial markets. **Figure 4** shows the estimated IRFs between the Realized Variance (COMRV) in the Commodity Market index and the logarithm of the Financial Uncertainty (FU) index, along with the 95% bootstrapped confidence intervals.

[Insert Figure 4 Here]

From **Figure 4** we can observe that an innovation in Financial Uncertainty (FU) results to an instantaneous increase in commodity price volatility. More specifically, a 1% positive shock in FU results to a persistent increase in commodity market volatility which reaches its maximum (0.4%) in the first 5 months after the initial FU shock and remains positive and statistically significant (within the bootstrapped confidence interval) for 14 months after the initial shock. Our results, thus, indicate that both the JNL MU and FU shocks have a significant (in terms of magnitude) and long-lasting impact in commodity market volatility. In other words, when the future state of the macroeconomy and the financial system becomes foggier, the price variability in commodity markets increases as a response. Our VAR analysis additionally shows that the MU shocks have a more significant impact on the RV in

commodity markets when compared to the respective impact of FU shocks. Our results show that the uncertainty about macroeconomic conditions seems to be the most important factor which drives time variation in commodity market volatility, when compared to the uncertainty about the conditions in financial system. In addition, **Figure 4** shows that the IRFs of FU to the commodity RV shocks are statistically insignificant. These empirical findings show that, while macroeconomic uncertainty is significantly reduced after the occurrence of large volatility swings in commodity markets, the financial uncertainty remains unaffected and immune to changes in commodity market turbulence. Unlike the MU index, the FU index does not have a significant response to commodity market volatility shocks. The Granger causality tests in **Table 2** lead us to the same conclusion, since we confirm a bi-directional causality between MU and Commodity market RV, and unidirectional causality from FU to Commodity Market RV.

3.2.2 The Impact of Observable Uncertainty Shocks

In this section we present the results of our VAR analysis when we use some widely accepted proxies for economic uncertainty which are based on observable variations in variables which are closely connected to macroeconomic fluctuations and uncertainty. For example, Bloom (2009) proposes the stock-market uncertainty index (VXO) and the volatility of the S&P 500 price index (SP500RV) as proxies for economic uncertainty. In addition, Baker *et al.* (2016) construct an Economic Policy Uncertainty index (EPU) which quantifies the economic policy uncertainty and it is based on newspaper articles. The analytical methodology for the construction of the EPU index and its respective components (EPU news Policy Uncertainty index (EPUNEWS), Fiscal Policy Uncertainty Index (EPUFISC) and Monetary Policy Uncertainty index (EPUMON)) can be found in Baker *et al.* (2016). Unlike the JNL MU and FU uncertainty series, the EPU index, the VXO index and the realized variance of the returns of the S&P 500 index (SP500RV), are observable indicators of economic fluctuations and they may fluctuate for reasons which are uncorrelated with economic uncertainty.¹⁸ Therefore, by estimating the impact of these alternative

¹⁸ Jurado *et al.* (2015) claim that “the stock-market volatility can change over time even if there is no change in uncertainty about economic fundamentals, if leverage changes, or if movements in risk aversion or sentiment are important drivers of asset market fluctuations. Cross sectional dispersion in the individual stock returns can fluctuate without any change in uncertainty if there is heterogeneity in the loadings of the common risk factors.” In addition, Bekaert *et al.* (2013) give further empirical

uncertainty measures, we can empirically examine which type of uncertainty shock matters most for commodity investors and producers, the unobservable or the observable measure. **Figure 5** shows the IRFs between the EPU index, as well as its components, and the RV in commodity market index.

[Insert Figure 5 Here]

Both the magnitude and the responses of the Realized Variance of the commodity price index to EPU shocks are much smaller when compared to the respective response of the commodity price RV to MU shocks that was presented in the previous section (**Figure 3**). For example, the commodity price index increases by 0.03% (3 basis points) in response to a positive 1% EPU shock. This effect is statistically significant only for the first month and vanishes after the second month. In addition, the response of commodity price volatility to the uncertainty about economic news (the news component of the Economic Policy Uncertainty index (EPUNEWS)) is of similar magnitude. These results show that, unlike the stock-market volatility (Antonakakis *et al.*, 2013; Liu and Zhang, 2015; Arouri *et al.*, 2016), the commodity market volatility seems to be relatively immune and less significantly affected by the observed uncertainty measures about future economic policy. Any kind of economic news which reveal a more uncertain economic environment, have a small and transitory impact on the volatility of commodity prices. In addition, the fiscal and the monetary policy components of the uncertainty index, have both a small negative impact on commodity price volatility. This negative impact of monetary (EPUMON) and fiscal policy uncertainty (EPUFISC) shocks vanishes after the second month of the initial shock.

Moreover, we estimate the baseline VAR model of Equation (2), in which we use some additional proxies of economic uncertainty, which have been proposed in the relevant literature, such as the stock-market volatility of the S&P 500 index, the VXO implied volatility index and the Financial Regulation index (FRU). **Figure 6** shows the estimated IRFs for the commodity RV-uncertainty pair when we use the

support to this claim since they show that the time varying volatility (the VIX index) in the equity market can be decomposed to investor's risk aversion and to economic uncertainty. This results show that the equity market volatility may change due to changes in risk aversion without any necessary change in the economic uncertainty.

RV of the S&P 500 index, the VXO index and the FRU index as alternative measures of uncertainty in the VAR model.

[Insert Figure 6 Here]

The stock-market volatility and financial regulation uncertainty shocks have positive, but small and transitory impact on commodity price volatility. For example, an one percentage point (100 basis points) shock in the logarithm of the VXO index increase the volatility in commodity prices by nine (0.09%) and two (0.02%) basis points respectively, with the effect being statistically insignificant. Overall, our results cannot verify the volatility spillovers hypothesis (Aroui *et al.*, 2011; Diebold and Yilmaz, 2012). While for example, Diebold and Yilmaz (2012) find that there are significant volatility spillover effects from equity to commodity markets, our VAR analysis shows that the impact of stock-market volatility is transitory and small.

3.2.3 The Impact of Uncertainty Shocks on the Volatility of Individual Commodity Markets

In this section, we conduct a disaggregated VAR analysis in which we measure the impact of uncertainty shocks on the Realized Variance (COMRV) of individual commodity prices. Thus, instead of measuring the impact of the broad commodity price index, we measure the impact of uncertainty shocks on the volatility of various agricultural, metals and energy commodities. By this approach, we implicitly examine whether economic uncertainty is a common volatility risk factor in, not only the aggregate, but also the individual commodity markets. We estimate 14 models of our baseline VAR of Equation (2), in which we use the RV of each one of the 14 individual commodity prices instead of the broad commodity index. We employ the VAR analysis using the MU and FU measures as economic uncertainty proxies, since we have shown in the previous sections that the RV of the commodity price index has an instant and highly persistent response to MU and FU shocks only.¹⁹ **Figure 7** shows the estimated IRFs for the VAR models in which the volatility of the various

¹⁹ For brevity, we do not report the responses of the volatility of commodity prices to EPU and stock-market volatility shocks like we did in subsection 3.2.2 for the broad commodity price index. The responses to EPU and stock-market volatility shocks of the volatility series of individual commodities are found to be insignificant. These results can be provided upon request.

agricultural commodity prices are used as the endogenous variable and MU as the economic uncertainty measure.

[Insert Figure 7 Here]

The estimated responses of the volatility of agricultural products on the MU shocks are all positive and statistically significant. Our VAR analysis shows that an 1% MU shock results to an approximately 0.2-0.5% increase in the monthly Realized Variance (RV) of the agricultural commodity futures markets. This effect is persistently positive and reaches its maximum 2-3 months after the initial shock in all agricultural commodities under consideration. The estimated response of the RV of corn, wheat and sugar prices to an MU shock is more persistent (the effect remains statistically significant for many months after the initial shock). On the other hand, the IRFs show that the MU series is relatively immune to volatility shocks of agricultural commodities. The only exemptions are sugar and wheat which have a positive and significant impact on MU. **Figure 8** shows the estimated IRFs between Macroeconomic Uncertainty (MU) and the Realized Variance in energy commodity markets.

[Insert Figure 8 Here]

The results indicate that the responses of the volatility of the energy commodity markets are more instant and long lasting when compared with agricultural markets. For example, a 1% MU shock results to a 2.5% increase in the volatility of crude oil price, and to a 2.2% increase in the volatility of heating oil price, with the effect being statistically significant from the first month until almost 18 months after the initial macroeconomic shock. In addition, a positive shock in the volatility of energy commodity markets reduces macroeconomic uncertainty. These results indicate a bi-directional causal relationship between energy commodity markets and macroeconomic uncertainty. These findings are in line with those of Hamilton (1983), Ferderer (1996), Hamilton (2003) and Elder and Serletis (2010) according to which oil price and volatility shocks are closely linked with the state of the macroeconomy and are significant indicators of U.S. economic recessions. Finally, **Figure 9** shows the estimated IRFs between MU and the commodity in metals commodity markets.

[Insert Figure 9 Here]

The estimated IRFs from **Figure 9** show that the effect of MU shocks in metals markets is again positive and statistically significant. More specifically, our estimated IRFs show that a 1% positive MU shock results to an almost equal magnitude increase (about 1%) on the volatility of the copper, silver and platinum prices, with the effect being statistically significant for about 15 months after the initial MU shock. The effect of MU shocks is slightly lower on the volatility of gold prices. We have to state here that the impact of Financial Uncertainty (FU) shocks to the volatility of individual commodity prices is again significant, but of relatively less magnitude compared to respective impact of MU shocks. The results of our VAR analysis, in which we use the FU instead of the MU as endogenous variable in the VAR model for the agricultural, energy and metals markets, can be found in our online Appendix.

3.3 OLS Regression Models

In order to show the robustness of our VAR analysis, we additionally employ OLS regression models on the impact of economic uncertainty on the volatility of commodity prices. The estimated coefficients from these regressions provide an additional evidence of the impact of uncertainty shocks on the volatility in commodity markets. Overall, our regression analysis provides further empirical support to our findings, according to which the uncertainty shocks that matter most for commodity investors are the unpredictable ones. Here, we measure the impact of uncertainty shocks on the volatility of commodities, by regressing the alternative uncertainty measures on the realized variance of the commodity futures S&P GSCI market index (COMRV). **Table 3** reports the results of these univariate regression models in which we use all the alternative, widely accepted, measures of economic uncertainty as explanatory variables.

[Insert Table 3 Here]

The results from **Table 3** indicate that all economic uncertainty measures have a positive and statistically significant impact on the realized variance of the commodity

price index, with the only exemption that of the fiscal policy uncertainty index. The insignificance of the fiscal policy uncertainty shocks is somewhat expected, since there is no empirical (or theoretical) evidence, linking commodity markets with fiscal policy. The coefficient of macroeconomic uncertainty (MU) is strongly positively significant and is much larger compared to the estimated coefficients of the other proxies of economic uncertainty. For example, the coefficient of the MU index is 0.359, while the estimated coefficients of the EPU and the VXO indices are 0.049 and 0.059 respectively. These results show the higher impact of MU on the realized variance of the commodity price index.²⁰ In addition, the adjusted R^2 value of the univariate regression model in the case of the MU index as the only determinant of RV of the commodity price index, reaches 39.9%, while the respective R^2 values for the rest univariate models that uses alternative economic uncertainty proxies are less than 22%. Moving further, we control for several macroeconomic and stock-market variables, that have been proposed in the literature as determinants of commodity markets, and show that the large positive impact of economic uncertainty on commodity price volatility remains robust to the inclusion of these variables on the left-hand side of our regression models.²¹ **Table 4** shows the estimated coefficients of the multivariate OLS regression models.

[Insert Table 4 Here]

From **Table 4** we observe that the estimated coefficients of MU and FU remain positive and statistically significant, while the coefficients of the fiscal and monetary policy uncertainty become insignificant. Moreover, we employ a forecasting exercise in which we use the MU, FU and EPU series as predictors for the volatility of the commodity price index (COMRV). The forecasting regression model is given in the Equation (4) below:

²⁰ We have to state here that the significantly larger coefficients of the MU index do not result from a different scale in the measurement of the alternative uncertainty measure. In order to have comparable results, we have multiplied the MU and the FU indices with 100 in order to be directly comparable and in line with the other economic uncertainty measures.

²¹ For example, Arouri *et al.* (2011) report significant volatility spillovers from equity to commodity markets, while Frankel and Hardouvelis (1985), Frankel (2008), Gilbert (2010) and Gubler and Hertweck (2013) identify the structural linkages between macroeconomic factors and commodity price and volatility dynamics.

$$COMRV_t = b_0 + b_1 MU_{t-k} + e_t. \quad (4)$$

In Equation (4), MU represents the JLN macroeconomic uncertainty, but we employ the same regression model using the FU and the EPU uncertainty measures for robustness purposes. **Table 5** reports the relevant results. For brevity, we report only the slope (b_1) coefficients.

[Insert Table 5 Here]

The OLS regression results indicate that the MU and FU series are statistically significant predictors of the commodity price volatility for both short and long-term forecasting horizons ranging from one up to twelve months ahead. Unlike the MU and the FU measures, the EPU index gives statistically significant forecasts only for short-term (up to 3 months) forecasting horizons. Our empirical findings show, for the first time, the predictive information content of economic uncertainty measures on the volatility of commodity prices. Our results on the predictive power of the MU and FU series remain robust to the inclusion of various traditional determinants of the commodity price volatility like the interest rates, the growth in the manufacturing Industrial Production Index and the volatility of the S&P 500 stock market index.²²

4. Robustness

In the online Appendix we conduct various robustness checks to supplement our empirical results, following the VAR analysis and the OLS regression models of the main paper.

4.1 VAR Models

The Appendix provides additional robustness to our VAR results, by estimating the Orthogonalized Impulse Response Function (OIRFs) using a Cholesky decomposition and show the responses using the OIRFs instead of the Generalized IRFs which we report in the main empirical section. Furthermore, we report the estimated OIRFs for alternative VAR orderings, following, for example, the VAR ordering of

²² These results which provide robustness to the forecasting performance of the MU and FU series can be found in the online Appendix.

Bloom (2009) and Jurado *et al.* (2015), and we find that our results remain robust to different VAR orderings. Moreover, motivated by the empirical studies on the significance of exchange rates (Gilbert, 1989; Chen *et al.*, 2010) and crude oil price shocks (Du *et al.*, 2011; Nazlioglu and Soytas, 2012; Shang *et al.*, 2016) for the price and the volatility path of commodities, we additionally control for the U.S. effective exchange rate and the oil price shocks in our VAR model. Lastly, we report the Granger causality tests between the uncertainty measures and the realized volatility of agricultural, energy and metals prices and find that there is a bi-directional causality between MU and the volatility in energy markets, while there is a unidirectional causality from MU to the volatility of agricultural and metals prices. Lastly, our VAR models and Granger causality reveal a significant impact of FU shocks on the volatility of agricultural, metals and energy futures markets.

4.2 OLS Regression Models

The Appendix contains various robustness checks to the OLS regression models of the main paper. More specifically, in order to provide further robustness to our OLS estimates which are presented in Section 3.2, we employ the same regression models in which we use the log-levels instead of the log-differences of the manufacturing IPI and the manufacturing Employment. In this way, we use exactly the same explanatory variables in the regression models as with the VAR system and we show that our empirical results remain unaltered, and are unaffected by the stationarity condition of these two variables ($I(0)/I(1)$). We, additionally, apply the same OLS regression models in which we use the JLN macroeconomic and financial uncertainty series which have 1, 3 and 12 month forecasting horizons (i.e., the MU1, MU3 and MU12 and the FU1, FU3 and FU12 respectively). By this exercise we empirically verify that both MU and FU series have a significant impact on the commodity price volatility irrespective of the forecasting horizon that has been used for their construction. Furthermore, we run the OLS regression models and we additionally control for the exchange rate and crude oil shocks (for consistency reasons, the left-hand side variables in these regressions are similar with those that are included as endogenous variables in our 8-factor VAR model). These results show that the impact of MU and FU shocks remains nearly the same under this regression specification with that of our baseline regression model presented in **Table 4**. In order to control for the presence of

sluggish (and not contemporaneous) reactions of commodity markets to uncertainty shocks, we use the same OLS regression models of the Subsection 3.3 in which here we include the lagged (one month before) explanatory variables. These results provide further evidence regarding the persistence of the impact of uncertainty shocks on commodity market volatility. These findings can alternatively be viewed as a first indication of the forecasting power of the JLN economic uncertainty measure on the commodity price volatility. Lastly, we empirically show that the impact of MU and FU shocks on the volatility of individual commodity prices is robust to the inclusion of additional macroeconomic factors (like the oil price shocks and the exchange rates) which are found in the literature that are directly linked to the commodity price volatility.

4.3 Uncertainty Shocks During the Financialization Era

In this section of the Appendix, we conduct the empirical analysis for the post-2000 period, during which the financialization of commodities (i.e., a large inflow of investment into commodity markets) has taken place. The financialization of commodity markets has led to structural changes in the nature as well as their information content. Since the early 1990s, the large inflow of funds and the increased presence of financial investors in commodity markets have transformed commodities into a separate asset class which has become more integrated to the rest of the financial markets (see Irwin and Sanders, 2012; Cheng and Xiong, 2013; Silvennoinen and Thorp, 2015; Basak and Pavlova, 2016). Tang and Xiong (2012) find that the financialization of commodity markets lies behind the increased (post-2008) correlation of oil prices and non-energy commodity prices, and the increased volatility of non-energy commodities around 2008. Motivated by the empirical findings that show an increased interdependence between stock-market and commodity market returns during the financialization period (Buyuksahin and Robe, 2014; Adams and Gluck, 2015), we empirically examine whether the financialization process has increases the structural linkages between uncertainty shocks and commodity market volatility.

Our empirical results indicate that the financialization in commodity markets has increased the interdependence and the sensitivity of commodity market volatility to uncertainty shocks. We show that the explanatory power and the significance of all

the economic uncertainty measures have tremendously increased during the post-2000 period. These results contradict with those of Karali and Power (2013) who find that the commodity specific factors dominate the macroeconomic factors when they are used to explain the volatility during the recent 2006-2009 period in the U.S. agricultural, energy and metals futures markets. On the contrary, our econometric analysis reveals that the common macroeconomic uncertainty factor explains a larger part of the time variation of the commodity price volatility in the post-financialization era, while this is not the case in the pre-2000 period. In order to provide robustness to our evidence regarding the significance of the financialization period, we run the same regressions for the pre-financialization (pre-2000) period. The results for the pre-2000 period provide robustness to our findings, since we show that most of the economic uncertainty measures turn from significant to insignificant during the pre-2000 period. These results can be found in the online Appendix.

5. Conclusions

This paper investigates the impact of uncertainty shocks on commodity price volatility. Overall, our results show that macroeconomic uncertainty increases volatility in commodity markets. Our analysis indicates that the rising degree of unpredictability in the macroeconomy, which is proxied by the latent uncertainty measure of Jurado *et al.* (2015), has the most significant and persistent impact on the volatility of commodity prices. On the other hand, the observable economic uncertainty measures have a rather transitory and less significant impact on the volatility of commodity prices. Our results suggest that the more unpredictable the future state of the macroeconomy becomes, the more volatile the prices of commodities will be.

The policy implication behind these findings is that the adoption of appropriate monetary and fiscal policies that can lead to the reduction of the unpredictability of macroeconomic fluctuations, will reduce also the variability of commodity markets. Commodity market turbulence does not seem to arise because of the per se macroeconomic and/or stock-market fluctuations. On the contrary, it is affected by the rising degree of unpredictability of these fluctuations. Any macroeconomic policy that can result in reducing this unpredictability (e.g., an adoption of a highly transparent monetary policy), will implicitly reduce the fluctuations as well as the

instability in commodity prices. Our VAR and OLS regression models indicate that the JNL macroeconomic uncertainty proxy, explains a large part of the time-varying volatility in commodity markets and, in addition, has a high predictive power when is used as a volatility predictor into the left-hand side of our volatility forecasting OLS regressions.

This paper examines the impact of economic uncertainty on the volatility of commodity markets through a unified framework employing both observable and unobservable macroeconomic and financial uncertainty shocks. One direction for future research would be the empirical examination of the predictive power of economic uncertainty on the volatility of commodity prices. We believe that it would be of interest to examine the predictive information content of macroeconomic uncertainty, when compared to the already empirically verified specific predictors of volatility in commodity markets (e.g., the inventory level and the option-implied volatility). This information would be very useful for commodity investors, producers and trade-policy makers.

References

- Adams, Z., and Z., Gluck (2015). “Financialization in commodity markets: A passing trend or the new normal?” *Journal of Banking and Finance*, 60, 93-111.
- Antonakakis, N., Chatziantoniou, I., and G., Filis (2013). “Dynamic co-movements of stock-market returns, implied volatility and policy uncertainty.” *Economics Letters*, 120, 87-92.
- Anzuini, A., Lombardi, M.J., and P., Pagano (2013). “The impact of monetary policy shocks on commodity prices.” *International Journal of Central Banking*, 9(3), 125-150.
- Arouri, M., Estay, C., Rault, C., and D., Roubaud (2016). “Economic policy uncertainty and stock markets: Long-run evidence from US.” *Finance Research Letters*, 18, 136-141.
- Arouri, M., Jouini, J., and D.K., Nguyen (2011). “Volatility spillovers between oil prices and stock sector returns: implications for portfolio management.” *Journal of International money and finance*, 30(7), 1387-1405.
- Baker, S.R., Bloom, N., and S.J., Davis (2016). “Measuring economic policy uncertainty.” *Quarterly Journal of Economics*, 131(4), 1593-1636.
- Basak, S., and A., Pavlova (2016). “A model of financialization of commodities.” *Journal of Finance*, 71(4), 1511-1556.

- Batten, J.A., Ciner, C., and B.M., Lucey (2010). "The macroeconomic determinants of volatility in precious metals markets." *Resources Policy*, 35(2), 65-71.
- Bekaert, G., Hoerova, M., and M., Lo Duca (2013). "Risk, uncertainty and monetary policy." *Journal of Monetary Economics*, 60 (7), 771-788.
- Bloom, N. (2009). "The impact of uncertainty shocks." *Econometrica*, 77(3), 623-685.
- Buyuksahin, B., and M.A., Robe (2014). "Speculators, commodities and cross-market linkages." *Journal of International Money and Finance*, 42, 38-70.
- Caldara, D., Fuentes-Albero, C., Gilchrist, S., and E., Zakrajsek (2016). "The macroeconomic impact of financial and uncertainty shocks." *European Economic Review*, 88, 185-207.
- Chen, Y.C., Rogoff, K.S., and B., Rossi (2010). "Can exchange rates forecast commodity prices?" *Quarterly Journal of Economics*, 125(3), 1145-1194.
- Cheng, I.H., and W., Xiong (2013). "The Financialization of commodity markets." Working Paper No.19642, NBER.
- Christensen, B.J., and N.R., Prabhala (1998). "The relation between implied and realized volatility." *Journal of Financial Economics*, 50(2), 125-150.
- Colombo, V. (2013). "Economic policy uncertainty in the US: Does it matter for the Euro area?" *Economics Letters*, 121, 39-42.
- Diebold, F.X., and K., Yilmaz (2012). "Better to give than to receive: Predictive directional measurement of volatility spillovers." *International Journal of Forecasting*, 28(1), 57-66.
- Du, X., Cindy, L. Y., and D.J., Hayes (2011). "Speculation and volatility spillover in the crude oil and agricultural commodity markets: A Bayesian analysis." *Energy Economics*, 33(3), 497-503.
- Elder, J. (2017). "Oil price volatility: Industrial production and special aggregates." *Macroeconomic Dynamics*, forthcoming.
- Elder, J., and A., Serletis (2010). "Oil price uncertainty." *Journal of Money, Credit and Banking*, 42(6), 1137-1159.
- Ferderer, J.P. (1996). "Oil price volatility and the macroeconomy." *Journal of Macroeconomics*, 18(1), 1-26.
- Frankel, J.A. (2008). "The effect of monetary policy on real commodity prices." In J.Y. Campbell: *Asset Prices and Monetary Policy*, University of Chicago Press, pp. 291-333.
- Frankel, J.A. (2013). "Effects of speculation and interest rates in a 'carry trade' model of commodity prices." Written for Understanding International Commodity Price Fluctuations, an International Conference organized by the IMF and Oxford University.
- Frankel, J.A., and G.A., Hardouvelis (1985). "Commodity prices, money surprises and Fed credibility." *Journal of Money, Credit and Banking*, 17 (4), 425-438.
- Frankel, J.A., and A.K., Rose (2010). "Determinants of agricultural and mineral commodity prices." In *Inflation in an era of Relative Price Shocks*. Reserve Bank of Australia.

- Fong, W.M., and K.H., See (2001). "Modelling the conditional volatility of commodity index futures as a regime switching process." *Journal of Applied Econometrics*, 16(2), 133-163.
- Gilbert, C.L. (1989). "The impact of exchange rates and developing country debt on commodity prices." *Economic Journal*, 99(397), 773-784.
- Gilbert, C.L. (2010). "How to understand high food prices." *Journal of Agricultural Economics*, 61(2), 398-425.
- Gordon, G., and K.G., Rouwenhorst (2006). "Facts and fantasies about commodity futures." *Financial Analysts Journal*, 62, 47-68.
- Gospodinov, N., and S., Ng (2013). "Commodity prices, convenience yields and inflation." *Review of Economics and Statistics*, 95(1), 206-219.
- Gubler, M., and M.S., Hertweck (2013). "Commodity price shocks and the business cycle: Structural evidence from the US." *Journal of International Money and Finance*, 37, 324-352.
- Hamilton, J. D. (1983). "Oil and the macroeconomy since World War II." *Journal of Political Economy*, 91(2), 228-248.
- Hamilton, J. D. (2003). "What is an oil shock?" *Journal of Econometrics*, 113(2), 363-398.
- Hammoudeh, S., Nguyen, D.K., and R.M., Sousa (2015). "US monetary policy and sectoral commodity prices." *Journal of International Money and Finance*, 57, 61-85.
- Hattori, M., Schrimpf, A., and V., Shushko (2016). "The response of tail risk perceptions to unconventional monetary policy." *American Economic Journal:Macroeconomics*, 8(2), 11-136.
- Henzel, S.R., and M., Rengel (2017). "Dimensions of macroeconomic uncertainty: A common factor analysis." *Economic Inquiry*, 55(2), 843-877.
- Irwin, S.H., and D.R., Sanders (2012). "Financialization and structural change in commodity futures markets." *Journal of Agricultural and Applied Economics*, 44(3), 371-396.
- Jo, S. (2014). "The effects of oil price uncertainty on global real economic activity." *Journal of Money, Credit and Banking*, 46(6), 1113-1135.
- Joets, M., Mignon, V., and T., Razafindrabe (2016). "Does the volatility of commodity prices reflect macroeconomic uncertainty?" Working paper 607, Banque de France.
- Jurado, K., Ludvigson, S.C., and S., Ng (2015). "Measuring uncertainty." *American Economic Review*, 105(3), 1177-1216.
- Kang, W., and R.A., Ratti (2014). "Oil shocks, policy uncertainty and stock-market returns." *Journal of International Financial Markets, Institutions and Money*, 26, 305-318.
- Karali, B., and G.J., Power (2013). "Short-and long-run determinants of commodity price volatility." *American Journal of Agricultural Economics*, 95(3), 724-738.
- Kelly, B., Pastor, L., and P., Veronesi (2016). "The price of political uncertainty: Theory and evidence from the option market." *Journal of Finance*, 71(5), 2417-2480.

- Kilian, L. (2008). "Exogenous oil supply shocks: How big are they and how much do they matter for the US economy?" *Review of Economics and Statistics*, 90(2), 216-240.
- Kilian, L. (2009). "Not all oil price shocks are alike: Disentangling demand and supply shocks in the crude oil market." *American Economic Review*, 99(3), 1053-1069.
- Koop, G., Pesaran, M.H, and S.M., Potter (1996). "Impulse response analysis in non-linear multivariate VAR models." *Journal of Econometrics*, 74(1), 119-147.
- Liu, L., and T., Zhang (2015). "Economic policy uncertainty and stock market volatility." *Finance Research Letters*, 15, 99-105.
- Meinen, P., and O., Roehle (2017). "On measuring uncertainty and its impact on investment: Cross-country evidence for the euro area." *European Economic Review*, 92, 161-179.
- Nazlioglu, S., and U., Soytas (2012). "Oil price, agricultural commodity prices, and the dollar: A panel cointegration and causality analysis." *Energy Economics*, 34(4), 1098-1104.
- Ng, V.K., and S.C., Pirrong (1994). "Fundamentals and volatility: Storage, spreads and the dynamics of metals prices." *Journal of Business*, 67(2), 203-230.
- Pastor, L., and P., Veronesi (2012). "Uncertainty about government policy and stock prices." *Journal of Finance*, 67(4), 1219-1264.
- Pastor, L., and P., Veronesi (2013). "Political uncertainty and risk premia." *Journal of Financial Economics*, 110 (3), 520-545.
- Pesaran, H.H., and Y., Shin (1998). "Generalized impulse response analysis in linear multivariate models." *Economics Letters*, 58(1), 17-29.
- Pindyck, R.S., and J.J., Rotemberg (1990). "Excess co-movement of commodity prices." *Economic Journal*, 100, 1173-1189.
- Rahman, S., and A., Serletis (2011). "The asymmetric effects of oil price shocks." *Macroeconomic Dynamics*, 15(S3), 437-471.
- Shang, H., Yuan, P., and L., Huang (2016). "Macroeconomic factors and the cross-section of commodity futures returns." *International Review of Economics and Finance*, 45, 316-332.
- Silvennoinen, A., and S., Thorp (2015). "Financialization, crisis and commodity correlation dynamics." *Journal of International Financial Markets, Institutions and Money*, 24, 42-65.
- Tang, K., and W., Xiong (2012). "Index investment and the financialization of commodities." *Financial Analysts Journal*, 68(5), 54-74.
- Van Robays, I. (2016). "Macroeconomic uncertainty and oil price volatility." *Oxford Bulletin of Economics and Statistics*, 78(5), 671-693.
- Wang, Z., Fausti, S.W., and B.A., Qasmi (2012). "Variance risk premiums and predictive power of alternative forward variances in the corn market." *Journal of Futures Markets*, 32(6), 587-608.
- Watugala, S.W. (2015). "Economic uncertainty and commodity futures volatility." Working Paper No. 15-14, Office of Financial Research.

Tables

Table 1: Descriptive Statistics and Unit Root Tests

The table shows the descriptive statistics along with unit root tests (the Augmented Dickey-Fuller test with lag-length based on Schwarz information criterion with a maximum of 12 lags and the Philips-Perron test with Newey-West automatic lag-length). EPU stands for the logarithm of the Economic Policy Uncertainty index of Baker *et al.* (2016). The EPUNews, EPUMON, EPUFISC and FRU series are the respective components of the EPU index representing the uncertainty about macroeconomic news, monetary policy, fiscal policy and financial regulation respectively. The MU1, MU3 and MU12 and the FU1, FU3 and FU12 series represent the logarithms of the latent JLN macroeconomic and financial uncertainty series which refer to the uncertainty over the 1-month, 3-month and 12-month horizon respectively. VXO is the logarithm of the monthly level of the VXO index, SP500 is the logarithm of the monthly level of the S&P 500 stock-price index, SP500RV is the monthly realized variance of the daily returns of the S&P 500 index, TERM is the difference between the constant maturity 10-year U.S. government bond yield and the 3-month U.S. Treasury Bill rate, EXCH is the logarithm of the U.S. Effective exchange rate, and COMRV is the realized variance of the daily returns of the S&P GSCI commodity futures price index. Lastly, MEMP and MIPI are the monthly levels of Industrial production and employment respectively, while OILP is the logarithm of oil price. The data have monthly frequency and cover the period from January 1985 until December 2016.

	Mean	Std. Dev.	Minimum	Maximum	Skewness	Kurtosis	ADF test (level)	ADF test (differences)	PP test (level)	PP test (differences)
EPU	4.639	0.282	4.047	5.502	0.320	2.544	-3.532***		-5.617***	
EPUNews	4.636	0.334	3.802	5.648	0.405	3.101	-8.000***		-7.744***	
EPUMON	4.384	0.577	2.808	6.011	0.054	2.730	-10.026***		-10.043***	
EPUFISC	4.478	0.565	3.138	5.925	0.298	2.454	-6.615***		-6.559***	
FRU	4.198	0.973	1.777	6.777	-0.033	2.879	-3.244**		-9.562***	
MU1	4.165	0.105	4.019	4.644	1.709	7.687	-3.549***		-2.784*	
MU3 (MU)	4.356	0.089	4.231	4.772	1.778	8.219	-3.066**		-2.718*	
MU12	4.513	0.046	4.448	4.728	1.805	8.363	-2.995**		-2.687*	
FU1	4.482	0.187	4.162	5.037	0.496	2.563	-3.297**		-2.808*	
FU3 (FU)	4.534	0.145	4.290	4.958	0.485	2.506	-3.261**		-2.759*	
FU12	4.587	0.052	4.505	4.731	0.455	2.305	-3.086**		-2.578*	
VXO	2.948	0.352	2.332	4.181	0.605	3.270	-4.281***		-3.917***	
SP500RV	0.032	0.066	0.002	0.828	8.131	84.414	-11.737***		-12.120***	
COMRV	0.041	0.050	0.003	0.497	4.511	31.326	-5.153***		-11.047***	
TERM	0.019	0.011	-0.005	0.038	-0.247	2.059	-3.372**		-3.209**	
SP500	6.683	0.688	5.191	7.714	-0.534	2.026	-1.619	-18.346***	-1.609	-18.341***
MEMP	9.625	0.158	9.346	9.801	-0.497	1.634	-0.900	-3.569***	-0.357	-7.467***
MIPI	4.429	0.220	3.999	4.701	-0.593	1.783	-1.795	-5.789***	-1.822	-17.572***
EXCH	4.489	0.137	4.235	4.969	0.647	3.953	-3.671***		-3.232**	
OILP	3.523	0.658	2.423	4.897	0.429	1.806	-1.716	-14.243***	-1.597	-13.743***

Table 2: Granger Causality Tests Between Economic Uncertainty Measures and Commodity Market Volatility (Baseline VAR(4) Model)

The table shows the results of the Granger causality tests between the various economic and financial uncertainty measures and the commodity market uncertainty. The tests refer to the multivariate 6-factor VAR mode with 4 lags with the following VAR ordering :[MIPI MEMP log(Uncertainty) TERM SP500 COMRV]. The Uncertainty index is the variable which represents our proxy for economic uncertainty. The null hypothesis is that the independent variable does not Granger cause the dependent variable. With * , ** and *** we denote the rejection of the null hypothesis of no causality at the 10%, 5% and 1% level respectively. Panels A, B and C show the estimated results for the Granger causality tests when we use alternative proxies for economic uncertainty. The time series sample covers the period from January 1985 until December 2016.

Panel A: Unobservable Economic Uncertainty

Dependent variable	Independent variable	chi-square	p-value
COMRV	MU	35.78***	0.000
COMRV	FU	11.03**	0.026
MU	COMRV	16.48***	0.002
FU	COMRV	3.00	0.56

Panel B: Observable Economic Policy Uncertainty

Dependent variable	Independent variable	chi-square	p-value
COMRV	EPU	8.23*	0.083
COMRV	EPUNews	10.13**	0.038
COMRV	EPUMON	7.77*	0.100
COMRV	EPUFISC	11.99**	0.017
EPU	COMRV	7.12	0.127
EPUNews	COMRV	9.92**	0.042
EPUMON	COMRV	4.34	0.362
EPUFISC	COMRV	4.70	0.319

Panel C: Additional Observable Proxies for Economic Uncertainty

Dependent variable	Independent variable	chi-square	p-value
COMRV	FRU	9.49**	0.050
COMRV	VXO	1.279	0.865
COMRV	SP500RV	8.69*	0.069
FRU	COMRV	1.57	0.814
VXO	COMRV	4.60	0.330
SP500RV	COMRV	1.24	0.871

Table 3: The Impact of Alternative Economic Policy Uncertainty Indexes on the Volatility of the Commodity Futures Price Index

The table shows the results of the OLS regression models that measure the impact of the various measures of economic uncertainty on the realized variance of the S&P GSCI commodity futures price index. EPU stands for the logarithm of the Economic Policy Uncertainty index of Baker *et al.* (2016). The variables EPUNEWS, EPUFISC, EPUMON and FRU are the respective components of the EPU index representing uncertainty about macroeconomic news, fiscal policy, monetary policy, and financial regulations respectively. The MU and FU series represent the logarithms of the unobserved (latent - with 3-month horizon) macroeconomic and financial uncertainty measures estimated by Jurado *et al.* (2015). The VXO is the logarithm of the VXO index, the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index. In the bivariate regression model we report only the slope coefficient. The baseline regression model is given in the following equation:

$$COMRV_t = a_0 + a_1 \log(\text{Uncertainty})_t + u_t$$

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
EPU	0.049** (2.24)								
EPUNEWS		0.047** (2.50)							
EPUMON			0.017* (1.88)						
EPUFISC				0.016 (1.56)					
FRU					0.016** (2.31)				
MU						0.359*** (5.27)			
FU							0.138** (2.55)		
VXO								0.059** (2.44)	
SP500RV									0.359* (1.86)
<i>N</i>	384	384	384	384	380	384	384	384	384
<i>RMSE</i>	0.048	0.048	0.049	0.049	0.048	0.039	0.046	0.046	0.044
<i>R</i> ²	0.077	0.097	0.036	0.032	0.099	0.399	0.156	0.169	0.219
<i>Adj. R</i> ²	0.074	0.094	0.033	0.030	0.097	0.398	0.154	0.167	0.217

t-statistics in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 4: The Impact of Alternative Economic Policy Uncertainty Indexes on the Volatility of the Commodity Futures Price Index when Controlling for Macroeconomic and Stock-Market Factors

The $\Delta SP500$ is the log-difference of the S&P500 stock-market index, the TERM is the term spread, $\Delta MEMP$ is the log-difference of the Manufacturing U.S. Employment and $\Delta MIPI$ is the log-difference of the Manufacturing U.S. Industrial Production Index. EPU stands for the logarithm of the Economic Policy Uncertainty. The variables EPUNEWS, EPUFISC, EPUMON and FRU are the respective components of the EPU index representing uncertainty about macroeconomic news, fiscal policy, monetary policy, and financial regulations respectively. The MU and FU series represent the logarithms of the JLN unobservable macroeconomic and financial uncertainty measures. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$\Delta SP500$	-0.165** (-1.98)	-0.153* (-1.82)	-0.174** (-1.97)	-0.177** (-2.04)	-0.166** (-2.00)	-0.088 (-1.62)	-0.138* (-1.81)	-0.116 (-1.60)	-0.006 (-0.12)
TERM	-0.296 (-1.03)	-0.187 (-0.69)	0.046 (0.18)	-0.074 (-0.25)	-0.197 (-0.72)	0.172 (0.84)	0.015 (0.06)	0.060 (0.23)	-0.049 (-0.24)
$\Delta MEMP$	-4.517*** (-4.61)	-4.341*** (-4.02)	-4.557*** (-4.14)	-4.557*** (-4.34)	-4.426*** (-4.48)	0.426 (0.44)	-3.165*** (-3.38)	-2.906*** (-2.84)	-3.596*** (-3.74)
$\Delta MIPI$	-1.823** (-2.26)	-1.860** (-2.30)	-2.062** (-2.48)	-2.019** (-2.43)	-1.812** (-2.31)	-1.644*** (-3.10)	-2.138*** (-2.77)	-2.223*** (-2.82)	-1.859*** (-2.92)
EPU	0.033*** (2.66)								
EPUNEWS		0.027*** (2.62)							
EPUMON			0.004 (0.55)						
EPUFISC				0.006 (0.90)					
FRU					0.009** (2.38)				
MU						0.304*** (5.61)			
FU							0.074*** (3.08)		
VXO								0.035*** (2.59)	
SP500RV									0.271* (1.96)
<i>N</i>	383	383	383	383	379	383	383	383	383
<i>RMSE</i>	0.041	0.041	0.042	0.042	0.042	0.038	0.041	0.041	0.039
<i>R</i> ²	0.320	0.321	0.296	0.298	0.321	0.439	0.328	0.341	0.385
<i>Adj. R</i> ²	0.311	0.312	0.286	0.289	0.312	0.431	0.319	0.332	0.377

t-statistics in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 5: Forecasting the Volatility of the Commodity Price Index using the Macroeconomic Uncertainty (MU), Financial Uncertainty (FU) and Economic Policy (EPU) Uncertainty Indexes, with forecasting horizon between 1 to 12 months ahead

	(1)	(2)	(3)	(4)	(5)	(6)
MU_t	0.359*** (5.27)					
MU_{t-1}		0.351*** (4.97)				
MU_{t-2}			0.335*** (4.60)			
MU_{t-3}				0.312*** (4.25)		
MU_{t-6}					0.247*** (3.47)	
MU_{t-12}						0.151** (2.51)
N	384	383	382	381	378	372
$RMSE$	0.039	0.040	0.041	0.042	0.045	0.049
R^2	0.399	0.382	0.348	0.302	0.190	0.071
$Adj. R^2$	0.398	0.380	0.346	0.300	0.188	0.069
	(1)	(2)	(3)	(4)	(5)	(6)
FU_t	0.138** (2.55)					
FU_{t-1}		0.137** (2.56)				
FU_{t-2}			0.134*** (2.60)			
FU_{t-3}				0.126*** (2.65)		
FU_{t-6}					0.099*** (2.80)	
FU_{t-12}						0.036** (2.03)
N	384	383	382	381	378	372
$RMSE$	0.046	0.046	0.046	0.047	0.048	0.050
R^2	0.156	0.155	0.148	0.132	0.082	0.011
$Adj. R^2$	0.154	0.153	0.146	0.130	0.080	0.008
	(1)	(2)	(3)	(4)	(5)	(6)
EPU_t	0.049** (2.24)					
EPU_{t-1}		0.039* (1.85)				
EPU_{t-2}			0.035* (1.77)			
EPU_{t-3}				0.020 (1.34)		
EPU_{t-6}					-0.000 (-0.02)	
EPU_{t-12}						-0.015* (-1.90)
N	384	383	382	381	378	372
$RMSE$	0.048	0.049	0.049	0.050	0.051	0.051
R^2	0.077	0.047	0.037	0.013	0.000	0.007
$Adj. R^2$	0.074	0.044	0.035	0.010	-0.003	0.004

Figures

Figure 1: Commodity Market Volatility and Macroeconomic Uncertainty

The figure shows the contemporaneous time series movements between Macroeconomic Uncertainty (MU) and the Realized Variance of the Commodity Price Index (COMRV).

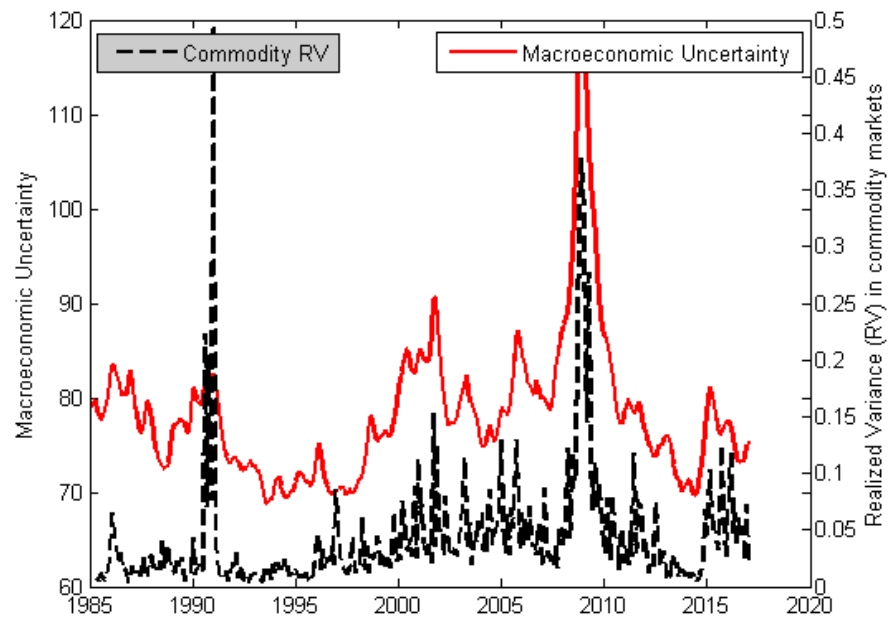


Figure 2: Commodity Market Volatility and Financial Uncertainty

This figure shows the contemporaneous time series movements between Financial Uncertainty (FU) and the Realized Variance of the Commodity Price Index (COMRV).

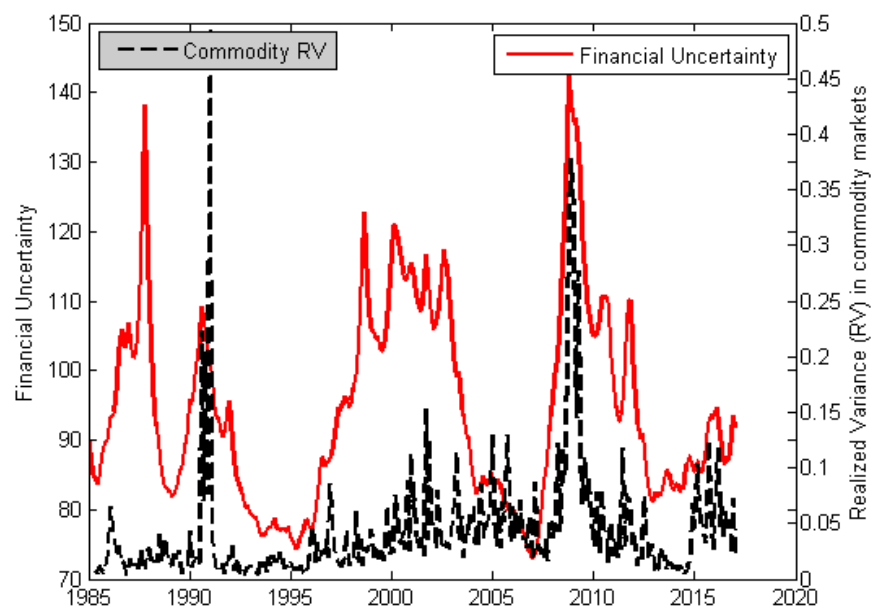


Figure 3: Impulse Response Functions (IRFs) between Commodity Market Volatility and Macroeconomic Uncertainty

The figure shows the estimated IRFs between the logarithm of Macroeconomic Uncertainty (MU) and the Realized Variance of the Commodity Futures Market Index (COMRV). The estimated responses are derived from the baseline 6-factor VAR model of Equation (2) and (3). The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated using 1000 replications.

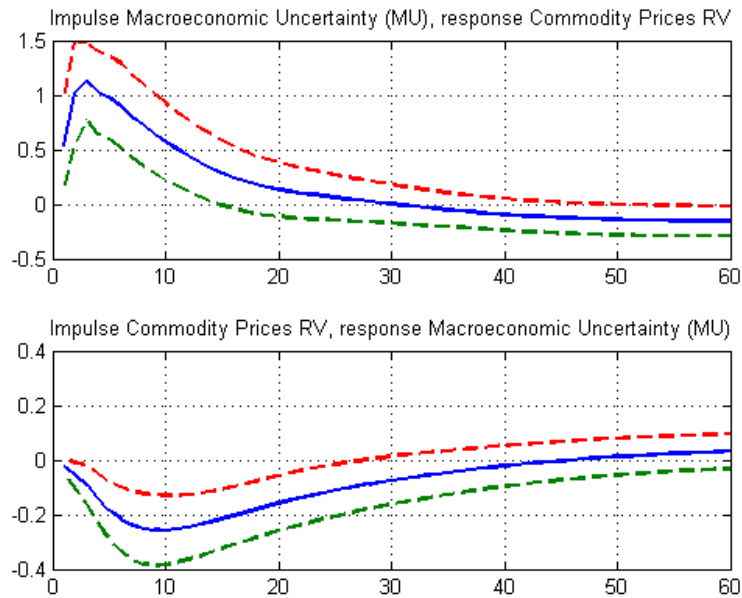


Figure 4: Impulse Response Functions (IRFs) between Commodity Market Volatility and Financial Uncertainty

The figure shows the estimated IRFs between the Financial Uncertainty (FU) and the Realized Variance of the Commodity Futures Market Index (COMRV). The estimated responses are derived from the baseline 6-factor VAR model of Equation (2) and (3). The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated using 1000 replications.

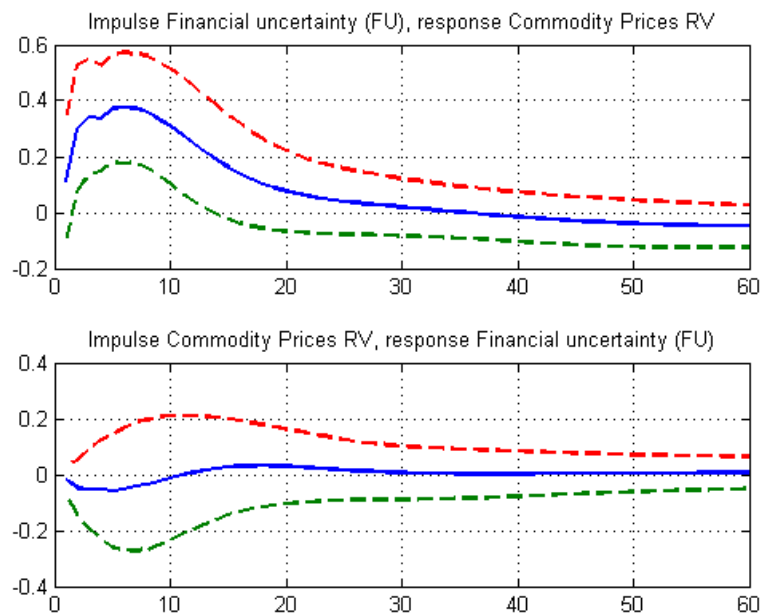


Figure 5: Impulse Response Functions (IRFs) between Commodity Market Volatility and Economic Policy Uncertainty

The figure shows the estimated IRFs between the various measures of economic policy uncertainty and the Realized Variance of the commodity futures price index (COMRV). The estimated responses are derived from the baseline 6-factor VAR model of Equation (2) and (3). Panel A shows the responses of Commodity Market Volatility shocks to Policy Uncertainty and Panel B shows the responses of Policy Uncertainty to Commodity Volatility shocks.

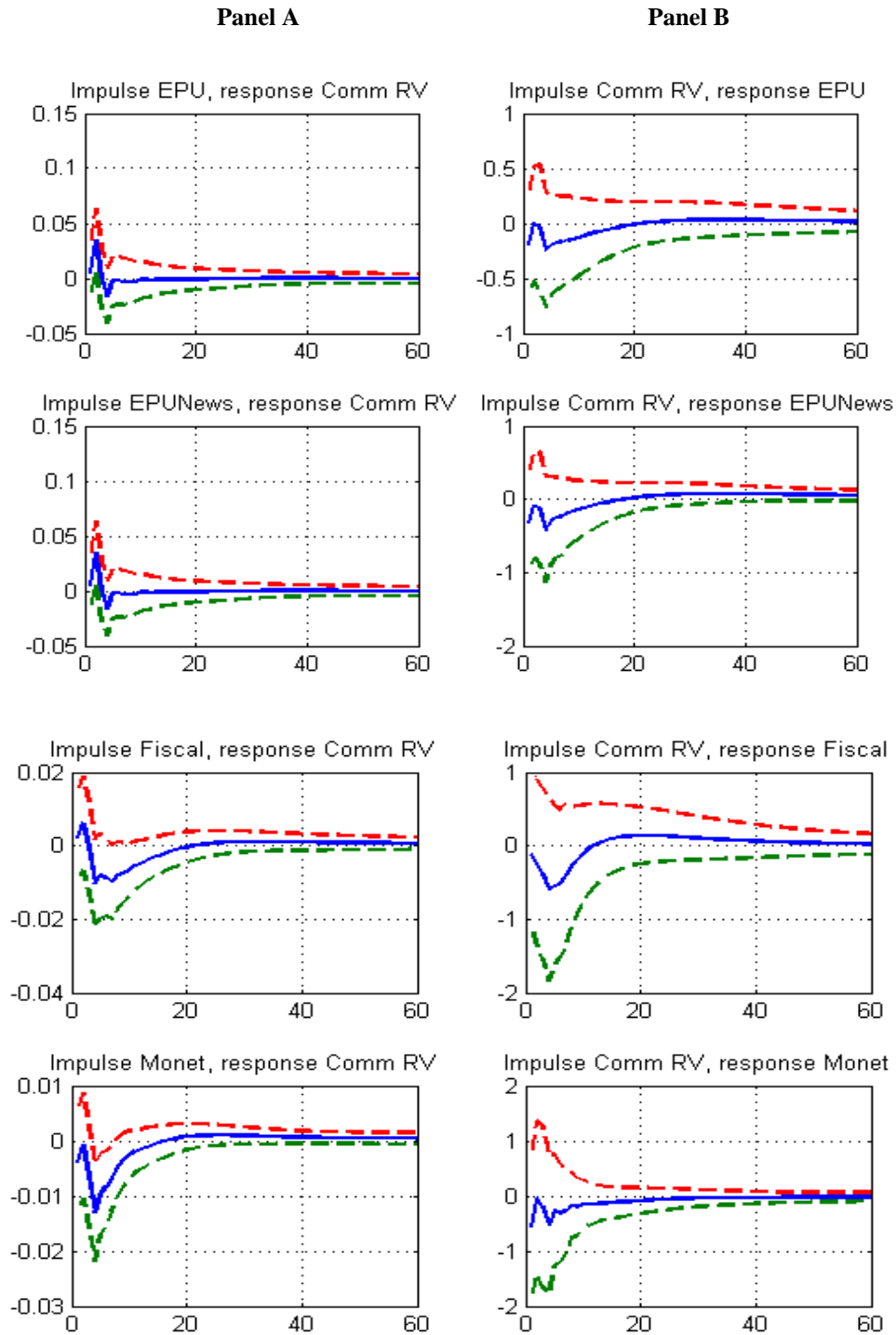


Figure 6: Impulse Response Functions (IRFs) between Commodity Market Uncertainty and Various Proxies of Financial Uncertainty

The figure shows the estimated IRFs between the various financial uncertainty measures and the Realized Variance of the commodity futures price index (COMRV). The estimated responses are derived from the baseline 6-factor VAR model of Equation (2) and (3). FRU stands for the logarithm of the Financial Regulation Uncertainty, SP500RV is the monthly realized variance of daily returns of the S&P 500 stock-market index, and VXO is the logarithm of the VXO index. Panel A shows the responses of COMRV to a shock in the logarithm of FRU, SP500RV and VXO series, respectively, while Panel B shows the responses of the logarithm of FRU, SP500RV and VXO to a shock in COMRV.

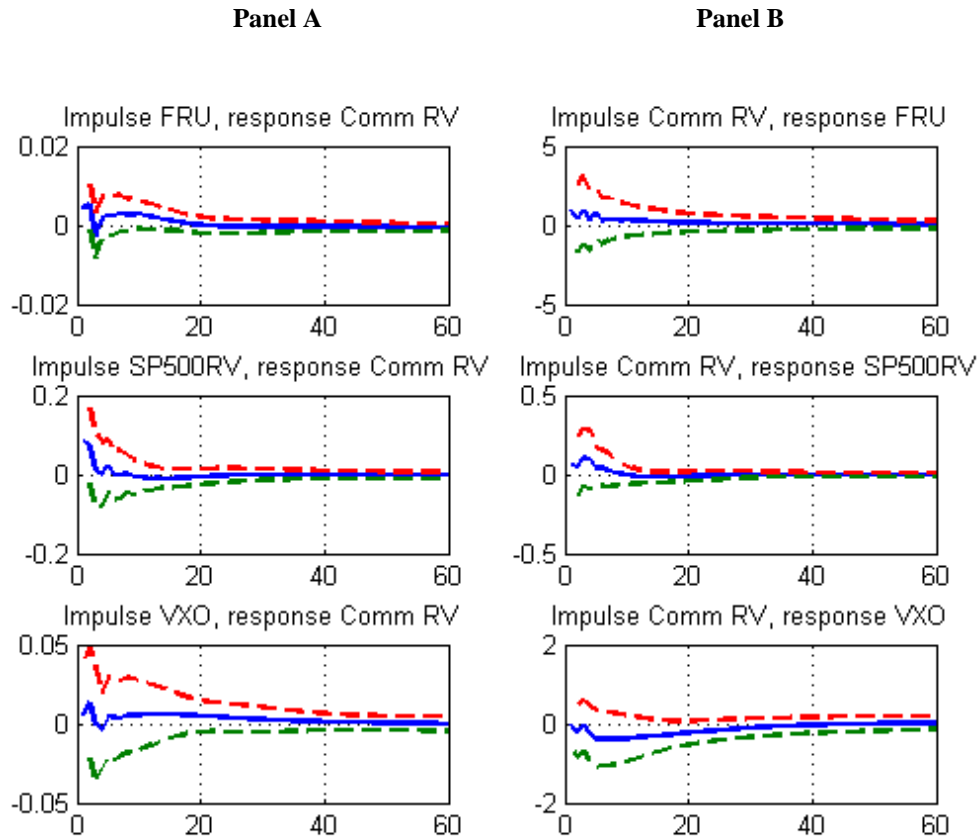


Figure 7: Impulse Response Functions (IRFs) between Realized Variance of Agricultural Commodity Markets and Macroeconomic Uncertainty

The figure shows the estimated IRFs between macroeconomic uncertainty (MU) and the Realized Variance (RV) of the agricultural commodity markets (cocoa, corn, cotton, soybeans, sugar, wheat). The estimated responses are derived from a baseline 6-factor VAR model provided in Equation (2) and (3). Panel A shows the responses of Agricultural Commodity Market Volatility to Macroeconomic Uncertainty shocks and Panel B shows the responses of Macroeconomic Uncertainty to Agricultural Commodity Volatility shocks.

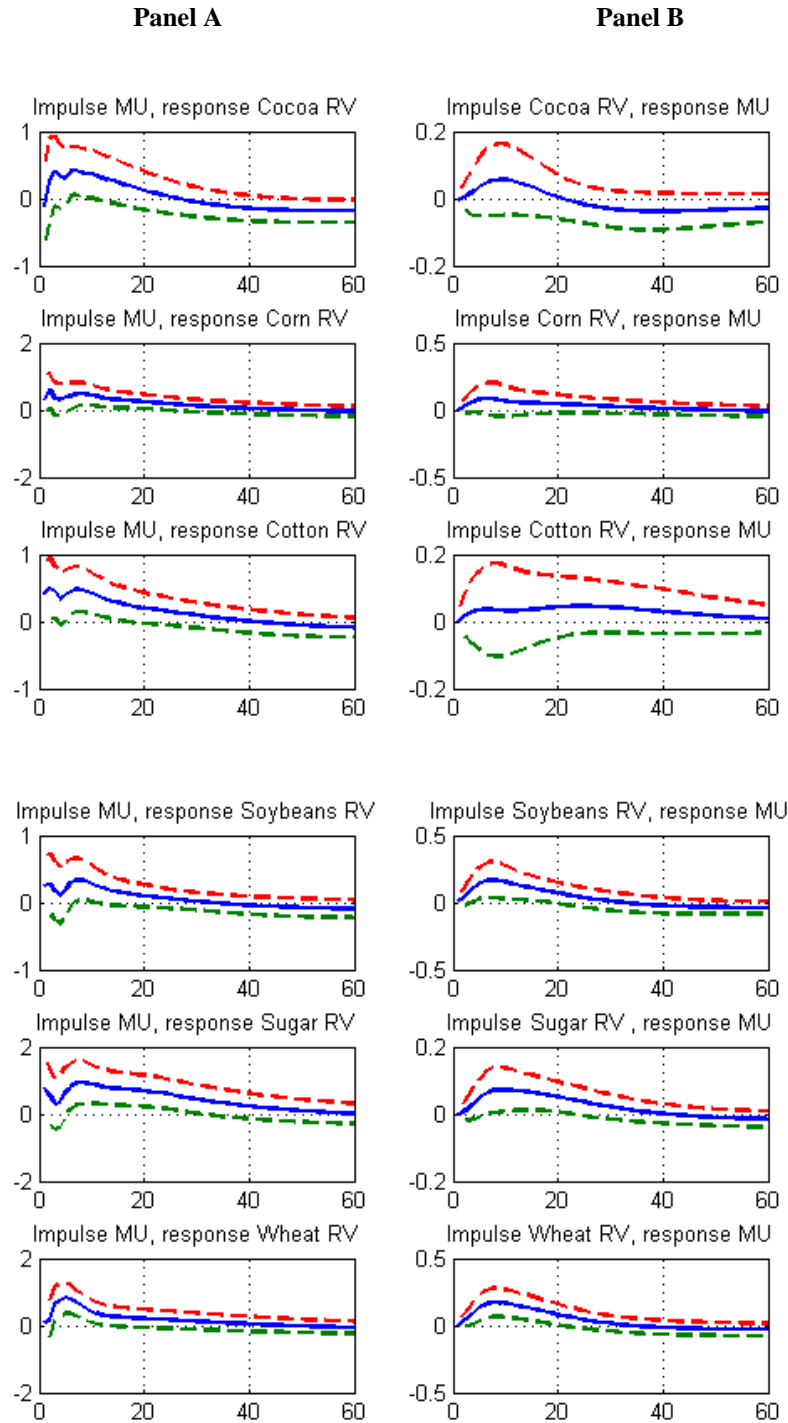


Figure 8: Impulse Response Functions (IRFs) between Realized Variance in Energy Commodity Markets and Macroeconomic Uncertainty

The figure shows the estimated IRFs between various measures of economic policy uncertainty and the Realized Variance (RV) of the energy commodity markets (crude oil, heating oil, petroleum and unleaded gasoline). The estimated responses are derived from a baseline 6-factor VAR model provided in Equation (2) and (3). Panel A shows the responses of Energy Commodity Market Volatility shocks to Macroeconomic Uncertainty shocks and Panel B shows the responses of Macroeconomic Uncertainty to Energy Commodity Volatility shocks.

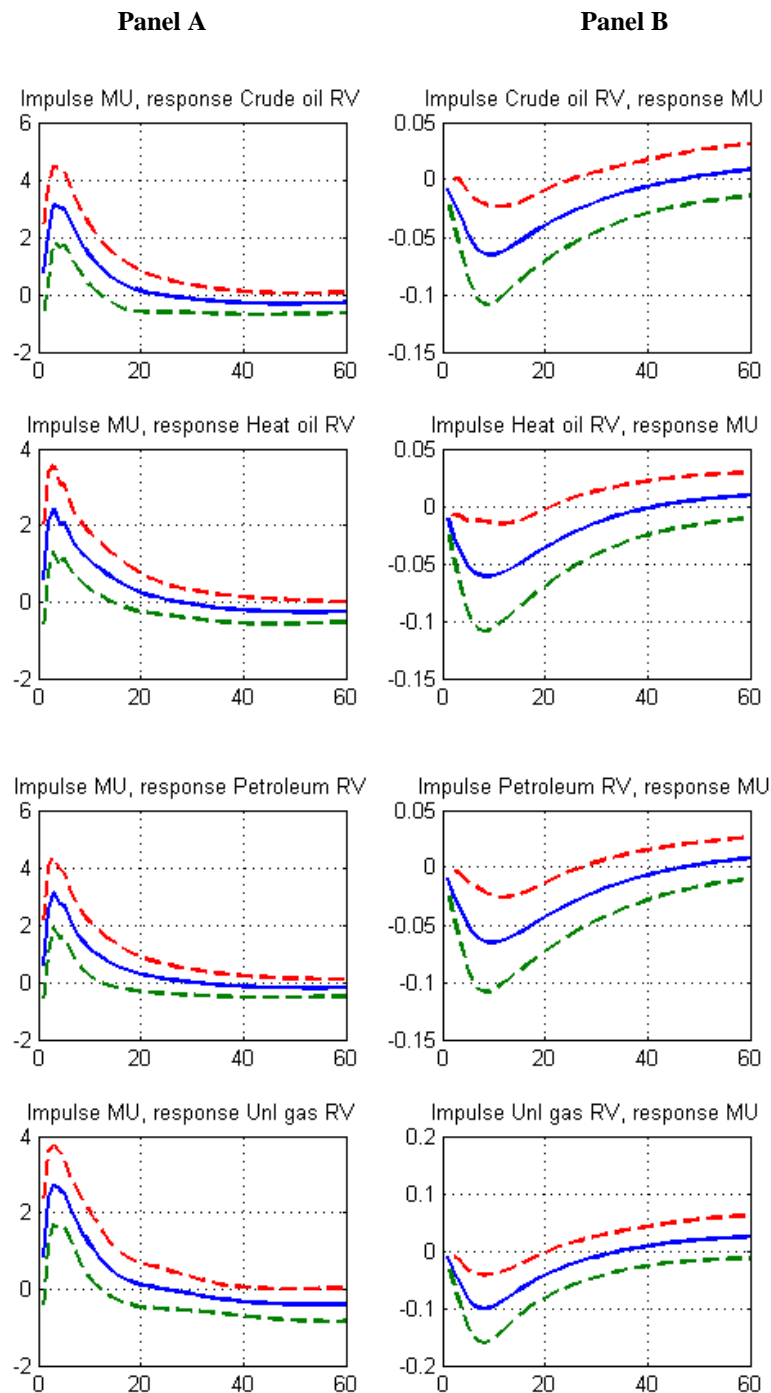
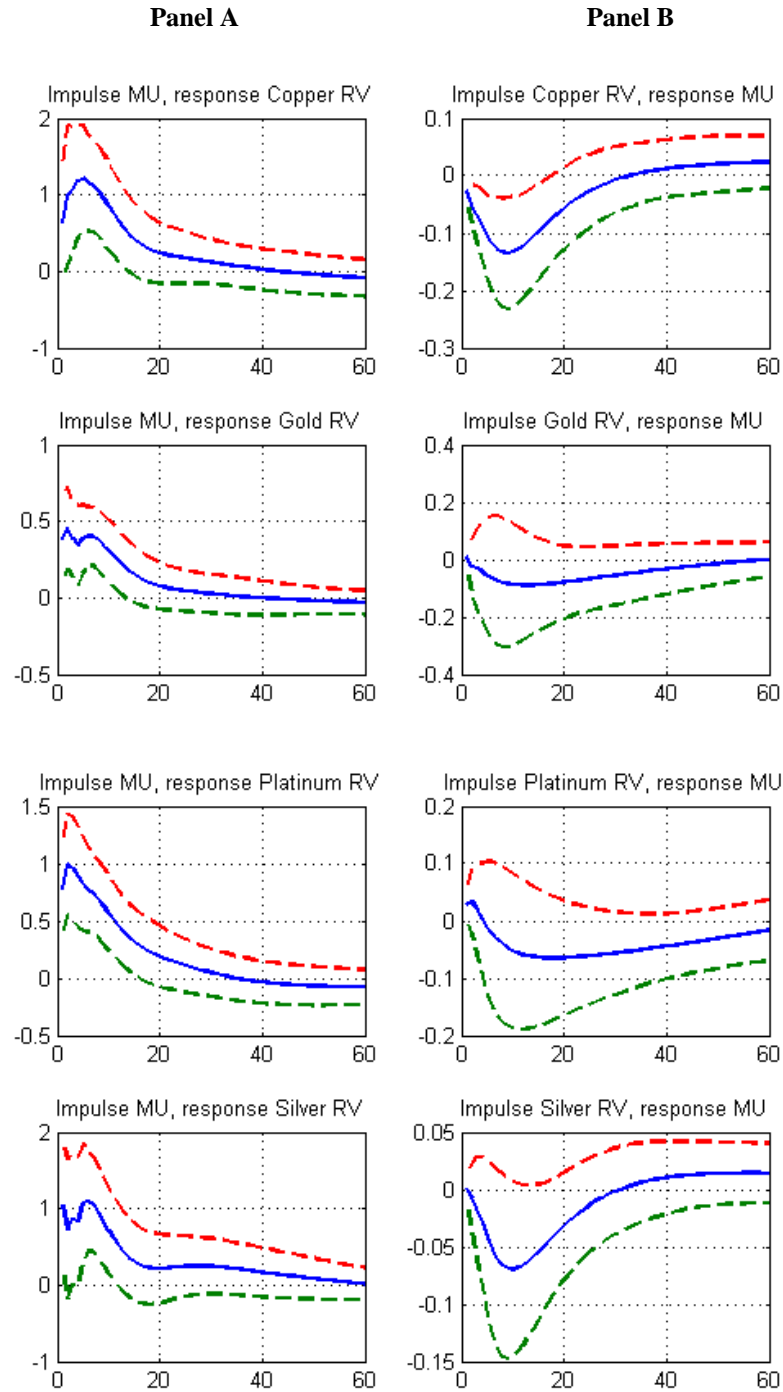


Figure 9: Impulse Response Functions (IRFs) between Realized Variance in Metals Commodity Markets and Macroeconomic Uncertainty

The figure shows the estimated IRFs between various measures of economic policy uncertainty and the Realized Variance (RV) in the metals commodity markets (gold, silver, copper and platinum). The estimated responses are derived from the baseline 6-factor VAR model provided in Equation (2) and (3). Panel A shows the responses of Metals Commodity Market Volatility shocks to Macroeconomic Uncertainty Shocks and Panel B shows the responses of Macroeconomic Uncertainty to Metals Commodity Market Volatility shocks.



The Impact of Uncertainty Shocks on the Volatility of Commodity Prices

Online Appendix

In this online Appendix we provide additional estimations and robustness checks on the OLS regressions and VAR models which are presented in the paper. **Section A** provides robustness to our OLS regression models, while **section B** provides robustness regarding the VAR estimates.

A. Robustness Tests for the OLS Regression Models

We provide robustness to our OLS regression models by estimating alternative specifications for the OLS regression equation. First, we estimate our baseline OLS regression model using the log-levels (instead of the log-differences) of the control variables which account for the macroeconomic and stock-market factors. In more details, we use the log-level of the manufacturing industrial production index (MIPI), the manufacturing employment (MEMP) and the monthly price of the S&P 500 stock-price index (SP500), instead of the respective log-differences which we have used in the baseline regressions in the main paper.

[Insert Table 1 Here]

The results from **Table 1** show that the estimated coefficients of the OLS regression model remain unaffected when we use the log-levels of the macroeconomic factors instead of the log-differences. We estimate an additional OLS regression model in which we control also for the exchange rates (EXCH) and oil price shocks (OILP). **Table 2** shows the OLS regression results when we use the log-level of the U.S. effective exchange rates and oil prices, while **Table 3** shows the respective results when we use the log-differences for these variables.

[Insert Table 2 Here]

[Insert Table 3 Here]

The OLS regression results from these alternative specifications reveal that our findings regarding the impact of macroeconomic (MU) and financial uncertainty (FU) shocks remain robust to the inclusion of the exchange rates and the oil prices (as well as to oil market returns).

Moreover, we provide robustness checks regarding the forecasting power of economic uncertainty on the volatility of the commodity price index by using the lagged values of the economic uncertainty and macroeconomic controls (such as the manufacturing industrial production index and the crude oil prices) in the left-hand side of the regression model. **Table 4** show the results when we use our lagged explanatory variables to forecasting the commodity price volatility having one-month forecasting horizon.

[Insert Table 4 Here]

The results from **Table 4** show the robustness of the forecasting power and the persistent impact of the economic uncertainty shocks on the volatility of commodity prices. More specifically, we find that the estimated coefficients of the various uncertainty measures remain positive and statistically significant when we regress these measures on the one-month ahead realized volatility of the commodity futures price index. The forecasting power of the economic uncertainty series remains robust to the inclusion of additional macroeconomic factors and the inclusion of the historical commodity price volatility.

Moreover, we strengthen our OLS regression results by estimating additional models in which we employ the alternative JLN macroeconomic and financial uncertainty series of Jurado *et al.* (2015) which have 1, 3 and 12 month forecasting horizons (MU1, MU3 and MU12 and FU1, FU3 and FU12 respectively). In the main paper we have used the MU3 (MU) and the FU3 (FU) uncertainty series (which have been constructed by Jurado *et al.* (2015) with a 3-month forecasting horizon), so, by this robustness test we empirically examine whether the MU and FU series are significant determinants of commodity price volatility irrespective of the forecasting horizon which have been used for their construction. **Table 5** reports the regression results of the univariate OLS regression models with the MU1, MU3, MU12 and FU1, FU3, FU12 uncertainty series as the explanatory variable, while **Tables 6** and **7** report the respective results of our multivariate OLS regression models in which we use the

MU and FU series with the different forecasting horizons and controlling also for the macroeconomic fundamentals.

[Insert Table 5 Here]

[Insert Table 6 Here]

[Insert Table 7 Here]

The results from **Tables 5, 6 and 7** show that the coefficients of MU1, MU3, MU12, FU1, FU3 and FU12 remain positive and statistically significant. This exercise empirically verifies that both MU and FU series have a significant impact on the commodity price volatility irrespective of the forecasting horizon which has been used for their construction. We additionally examine the impact of the economic uncertainty shocks on the volatility of commodity prices over the financialization (post-2000) and the pre-financialization (pre-2000) periods. **Tables 8 and 9** show the regression results for the baseline multivariate regression model over the post-financialization and the pre-financialization period respectively.

[Insert Table 8 Here]

[Insert Table 9 Here]

From **Table 8** we observe that the estimated coefficients of MU and FU are positive and statistically significant and that the adjusted R^2 values are much larger for the financialization (post-2000) period when compared with the adjusted R^2 values of the baseline regression model which are presented in the main paper (that refers to the full sample (1985-2016)). Furthermore, the estimated coefficients of the uncertainty shocks for the pre-financialization period become statistically insignificant (**Table 9**). These results reveal the fact that the financialization of commodity markets which has taken place in the last two-decades is a key factor behind the significant increasing impact of uncertainty shocks on commodity price volatility. Lastly, we estimate various OLS regression models in which we use, instead of the volatility of the commodity price index (COMRV), the volatility of the returns of agricultural, metals and energy commodity futures. **Tables 10, 11 and 12** report the regression results, using the EPU, the MU and the FU indices as explanatory variables respectively.

[Insert Table 10 Here]

[Insert Table 11 Here]

[Insert Table 12 Here]

From **Tables 10, 11 and 12** we can observe that the estimated regression coefficients of the MU and FU proxies are much higher and strongly significant positive when compared to the regression coefficients of the EPU index. These results provide further empirical support to our initial VAR estimates and conclusions, according to which the unobservable MU and FU shocks are those which have the most significant impact for commodity markets. We additionally observe that the impact of macroeconomic shocks is larger for the energy commodity volatility compared to the volatility of the agricultural and metals commodity markets. We additionally estimate the same regression models on individual commodity prices shown in **Tables 10, 11 and 12** in which we control also for additional macroeconomic and financial factors. **Tables 13, 14 and 15** show the respective OLS regression results in which we measure the impact of MU, FU and EPU shocks on the volatility of agricultural, metals and energy commodity futures respectively.

[Insert Table 13 Here]

[Insert Table 14 Here]

[Insert Table 15 Here]

The estimated regression results from **Tables 13, 14 and 15** show that the impact of MU and FU on the volatility of agricultural, metals and energy prices is robust to the inclusion of macroeconomic and stock-market factors.

B. Robustness Tests for the VAR Models

In this section we provide additional robustness checks for our VAR analysis. First, we estimate the orthogonalized impulse response functions for our baseline regression model given in Equation (3) of the main paper. **Figures 1 and 2** below show the estimated OIRFs when we use the MU and the FU as the endogenous variable at our baseline 6-factor VAR model.

[Insert Figure 1 Here]

[Insert Figure 2 Here]

The estimated OIRFs show that the responses of commodity market volatility to orthogonalized shocks in economic uncertainty are of same sign and magnitude as the Generalized IRFs which are estimated and reported in the main paper. Furthermore, in order to provide robustness to our VAR estimates, we estimate a VAR model with the same endogenous variables but with a VAR ordering similar to the VAR model of Jurado *et al.* (2015) and Baker *et al.* (2016). Our alternative VAR ordering is given in Equation (1) below:

$$Y_t = [SP500_t, MU_t, COMRV_t, TERM_t, MEMP_t, MIPI_t] \quad (1)$$

The estimated OIRFs for the VAR model with the ordering given in Equation (1) when using MU and FU factors are given in **Figures 3 and 4**.

[Insert Figure 3 Here]

[Insert Figure 4 Here]

The results shown in **Figure 3 and 4** clearly show that the estimated OIRFs remain unaltered and they are insensitive to the selection of the ordering of the variables in the VAR model. Overall, these results provide further robustness to the findings and conclusions which are presented in the Subsection 3.2 of the main paper. Moreover, we run an 8-factor VAR model in which we additionally include as endogenous variables the logarithm of the crude oil price and the logarithm of the U.S. effective exchange rate. The VAR ordering for the 8-factor VAR model is given in Equation (2):

$$Y_t = [MIPI_t, MEMP_t, MU_t, EXCH_t, TERM_t, SP500_t, OILP_t, COMRV_t] \quad (2)$$

Figures 5 and 6 show the estimated IRFs between the Realized Variance in the commodity futures price index (COMRV) and the MU and FU measures respectively. Under this 8-factor VAR specification, our basic findings and conclusions remain

unaltered. Using this VAR identification scheme, we again find that a MU shock has an equal in magnitude impact in the Realized Variance of the commodity price index (COMRV), when compared to that from the 6-factor VAR.

[Insert Figure 5 Here]

[Insert Figure 6 Here]

The estimated IRFs in **Figure 5** and **6** clearly indicate that our empirical results remain robust to the inclusion of these additional determinants of the price and volatility dynamics of commodity markets, like the exchange rate and the oil price shocks. From **Figure 5** we observe that a positive 1% shock in the MU increases the commodity price RV by 1.1% after 3 months, and a 1% positive shock in the FU increases the commodity price RV by approximately 0.25% after 3 months. These responses remain positive and statistically significant for about 15 months after the initial uncertainty shock. Lastly, the response of MU to RV shocks in the commodity price index is persistently negative, thus, our basic conclusion, according to which large volatility episodes in commodity markets are significant components of the heightened macroeconomic uncertainty (since the MU reduces right after the occurrence of a large volatility shock), remains valid under the 8-factor VAR identification scheme.

We continue by providing some additional estimates of the findings presented in Subsection 3.2.3 of the main paper and explore the impact of macroeconomic (MU) and financial (FU) uncertainty shocks on the volatility of agricultural, energy and metals commodity markets. **Tables 16** and **17** report the Granger causality tests for the uncertainty-commodity price Realized Variance pairs when we use the RV of the agricultural (cocoa, corn, cotton, wheat, soybeans, sugar), energy (crude oil, heating oil, unleaded gasoline, petroleum) and metals (copper, gold, silver, platinum) commodity futures prices. **Table 16** reports the results of the Granger causality tests when we use MU as the economic uncertainty measure in the VAR, while **Table 17** reports the respective results when we use FU as the economic uncertainty measure.

[Insert Table 16 Here]

[Insert Table 17 Here]

The Granger causality tests for the individual commodity volatility series indicate that there is a bi-directional causal relationship between volatility in energy commodity prices and macroeconomic uncertainty. In addition, we find strong causal relationships from FU to the RV of individual commodity price series. Unlike, MU, we do not find any bi-directional causal relationship between FU and the RV in energy commodity markets. This leads to the conclusion that, there is not any significant uncertainty transmission mechanism from the commodity markets to the rest of the financial markets. Lastly, **Figures 7, 8 and 9** show the estimated IRFs between FU and the volatility of agricultural, energy and metals commodity futures respectively.

[Insert Figure 7 Here]

[Insert Figure 8 Here]

[Insert Figure 9 Here]

The estimated IRFs in **Figures 7, 8 and 9** indicate that the FU shocks have a nearly identical impact (in terms of persistence) but slightly less in terms of magnitude (when compared with the impact of MU shocks) on the volatility of agricultural, metals and energy commodity futures prices.

Table 1: Impact of alternative Economic Uncertainty Measures on commodity price volatility when controlling for the log-level of macroeconomic factors.

In this table we report the results of our baseline multivariate regression model in which we use the log-levels of the Manufacturing Industrial Production (MIPI), the Manufacturing Employment (MEMP) and the S&P500 stock-market index (SP500). EPU stands for the Economic Policy Uncertainty. The variables EPUNews, EPUFISC, EPUMON and the Financial Regulation Uncertainty (FRU) are the respective components of the Baker *et al.* (2016) EPU index representing the Uncertainty about macroeconomic news, about fiscal policy, monetary policy, and financial regulation uncertainty respectively. MU and FU are the logarithms of the unobservable macroeconomic and financial uncertainty. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
SP500	-0.054* (-1.94)	-0.065** (-2.17)	-0.050* (-1.65)	-0.056* (-1.88)	-0.048* (-1.87)	0.009 (0.70)	-0.063** (-2.53)	-0.039* (-1.91)	-0.038** (-2.33)
TERM	-0.159 (-0.51)	-0.240 (-0.75)	-0.125 (-0.38)	-0.210 (-0.58)	-0.104 (-0.33)	0.296 (1.10)	-0.353 (-1.24)	-0.328 (-1.09)	-0.264 (-0.96)
MEMP	0.027 (0.77)	-0.025 (-0.73)	-0.070* (-1.74)	-0.031 (-0.89)	0.014 (0.37)	0.009 (0.39)	-0.092** (-2.37)	-0.105** (-2.51)	-0.068** (-2.05)
MIPI	0.239** (2.52)	0.230** (2.44)	0.182* (1.93)	0.218** (2.28)	0.207** (2.33)	0.003 (0.07)	0.186** (2.54)	0.115* (1.73)	0.126** (2.30)
EPU	0.063*** (2.69)								
EPUNews		0.045** (2.50)							
EPUMON			0.020** (2.36)						
EPUFISC				0.019* (1.82)					
FRU					0.017** (2.37)				
MU						0.353*** (4.78)			
FU							0.134*** (2.99)		
VXO								0.060*** (2.87)	
SP500RV									0.328* (1.85)
<i>N</i>	384	384	384	384	380	384	384	384	384
<i>RMSE</i>	0.045	0.045	0.046	0.046	0.046	0.039	0.043	0.043	0.042
<i>R</i> ²	0.192	0.191	0.162	0.149	0.187	0.413	0.259	0.274	0.287
<i>Adj. R</i> ²	0.182	0.180	0.151	0.138	0.177	0.405	0.249	0.265	0.278

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 2: Impact of alternative Economic Uncertainty Measures on commodity price volatility when controlling for the log-level of macroeconomic factors, oil prices and exchange rates.

In this table we report the results of our baseline multivariate regression model in which we use the log-levels of macroeconomic factors, oil price shocks and exchange rates. More specifically, we control for the Manufacturing Industrial Production (MIPI), the Manufacturing Employment (MEMP) the S&P500 stock-market index (SP500), the oil price (OILP) and the US exchange rate (EXCH). EPU stands for the Economic Policy Uncertainty. The variables EPUNEWS, EPUFISC, EPUMON and the Financial Regulation Uncertainty (FRU) are the respective components of the Baker *et al.* (2016) EPU index representing the Uncertainty about macroeconomic news, about fiscal policy, monetary policy, and financial regulation uncertainty respectively. MU and FU are the logarithms of the unobservable macroeconomic and financial uncertainty. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
SP500	-0.060* (-1.79)	-0.068** (-1.98)	-0.058 (-1.58)	-0.063* (-1.79)	-0.057* (-1.80)	0.005 (0.33)	-0.067** (-2.47)	-0.037 (-1.55)	-0.044** (-2.06)
TERM	-0.197 (-0.68)	-0.251 (-0.88)	-0.201 (-0.66)	-0.286 (-0.87)	-0.206 (-0.69)	0.312 (1.26)	-0.357 (-1.38)	-0.276 (-1.02)	-0.322 (-1.31)
MIPI	0.000 (0.00)	-0.039 (-0.40)	-0.114 (-1.21)	-0.074 (-0.74)	-0.038 (-0.42)	-0.021 (-0.35)	-0.107 (-1.49)	-0.087 (-1.25)	-0.099 (-1.28)
MEMP	0.258** (2.36)	0.242** (2.21)	0.206* (1.80)	0.240** (2.13)	0.234** (2.20)	0.021 (0.39)	0.200** (2.37)	0.110 (1.35)	0.142* (1.96)
EXCH	-0.000 (-0.64)	-0.000 (-0.61)	-0.000 (-0.40)	-0.000 (-0.17)	-0.000 (-0.28)	-0.001** (-2.42)	-0.000 (-1.23)	-0.000 (-0.81)	-0.000 (-0.10)
OILP	-0.010 (-0.61)	-0.007 (-0.39)	-0.013 (-0.75)	-0.012 (-0.66)	-0.015 (-0.91)	-0.021* (-1.85)	-0.010 (-0.73)	0.001 (0.07)	-0.008 (-0.55)
EPU	0.064** (2.46)								
EPUNEWS		0.045** (2.25)							
EPUMON			0.020** (2.18)						
EPUFISC				0.019* (1.69)					
FRU					0.017** (2.32)				
MU						0.380*** (5.26)			
FU							0.139*** (2.95)		
VXO								0.061*** (2.77)	
SP500RV									0.325* (1.85)
<i>N</i>	384	384	384	384	380	384	384	384	384
<i>RMSE</i>	0.045	0.045	0.046	0.046	0.045	0.038	0.043	0.043	0.042
<i>R</i> ²	0.197	0.193	0.167	0.153	0.194	0.443	0.268	0.277	0.289
<i>Adj. R</i> ²	0.182	0.178	0.152	0.137	0.178	0.432	0.254	0.263	0.276

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 3: Impact of alternative Economic Uncertainty Measures on commodity price volatility when controlling for the log-difference of macroeconomic factors, oil prices and exchange rates.

In this table we report the results of our baseline multivariate regression model in which we use the log-differences of macroeconomic factors, oil price shocks and exchange rates. More specifically, we control for the log-difference of the Manufacturing Industrial Production (MIPI), the Manufacturing Employment (MEMP) the S&P500 stock-market index (SP500), the oil price (OILP) and the US exchange rate (EXCH). EPU stands for the Economic Policy Uncertainty. The variables EPUNEWS, EPUFISC, EPUMON and the Financial Regulation Uncertainty (FRU) are the respective components of the Baker *et al.* (2016) EPU index representing the Uncertainty about macroeconomic news, about fiscal policy, monetary policy, and financial regulation uncertainty respectively. MU and FU are the logarithms of the unobservable macroeconomic and financial uncertainty. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$\Delta SP500$	-0.154** (-2.08)	-0.145* (-1.92)	-0.164** (-2.11)	-0.167** (-2.16)	-0.156** (-2.11)	-0.081 (-1.47)	-0.128* (-1.84)	-0.107 (-1.59)	-0.010 (-0.16)
TERM	-0.270 (-0.91)	-0.164 (-0.59)	0.046 (0.17)	-0.052 (-0.17)	-0.187 (-0.67)	0.167 (0.86)	0.014 (0.06)	0.062 (0.24)	-0.045 (-0.22)
$\Delta MEMP$	-4.669*** (-4.45)	-4.503*** (-3.90)	-4.710*** (-4.01)	-4.715*** (-4.19)	-4.579*** (-4.33)	0.188 (0.21)	-3.340*** (-3.36)	-3.126*** (-2.92)	-3.750*** (-3.73)
$\Delta MIPI$	-1.685** (-2.44)	-1.728** (-2.47)	-1.892*** (-2.69)	-1.863*** (-2.64)	-1.661** (-2.48)	-1.512*** (-3.17)	-1.969*** (-2.99)	-2.065*** (-3.01)	-1.755*** (-2.99)
$\Delta EXCH$	0.101 (0.77)	0.081 (0.62)	0.072 (0.56)	0.069 (0.53)	0.075 (0.55)	0.061 (0.49)	0.064 (0.48)	0.095 (0.69)	0.036 (0.30)
$\Delta OILP$	-0.071 (-1.46)	-0.071 (-1.45)	-0.081 (-1.64)	-0.079 (-1.60)	-0.075 (-1.60)	-0.066* (-1.68)	-0.079* (-1.72)	-0.069 (-1.60)	-0.055 (-1.54)
EPU	0.031** (2.43)								
EPUNEWS		0.025** (2.36)							
EPUMON			0.003 (0.50)						
EPUFISC				0.005 (0.73)					
FRU					0.008** (2.30)				
MU						0.297*** (6.38)			
FU							0.072*** (3.19)		
VXO								0.034*** (2.63)	
SP500RV									0.256* (1.95)
<i>N</i>	383	383	383	383	379	383	383	383	383
<i>RMSE</i>	0.041	0.041	0.042	0.042	0.041	0.037	0.041	0.040	0.039
<i>R</i> ²	0.338	0.338	0.317	0.319	0.340	0.453	0.348	0.358	0.395
<i>Adj. R</i> ²	0.326	0.325	0.305	0.306	0.328	0.443	0.336	0.346	0.384

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 4: The predictive power of the alternative economic uncertainty measures on commodity price volatility when controlling for the lagged macroeconomic factors, exchange rates and oil shocks.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
LCOMRV	0.401*** (2.76)	0.406*** (2.82)	0.418*** (2.90)	0.413*** (2.86)	0.388*** (2.63)	0.251** (2.40)	0.383*** (2.78)	0.379*** (2.74)	0.361*** (3.02)
LΔSP500	-0.134** (-2.20)	-0.119** (-1.97)	-0.134** (-2.15)	-0.148** (-2.40)	-0.143** (-2.31)	-0.101** (-2.03)	-0.132** (-2.24)	-0.105** (-1.99)	-0.080* (-1.79)
LTERM	-0.309 (-1.22)	-0.216 (-0.97)	-0.059 (-0.32)	-0.170 (-0.68)	-0.262 (-1.17)	0.127 (0.75)	-0.048 (-0.27)	-0.011 (-0.06)	-0.106 (-0.67)
LΔMEMP	-2.235*** (-3.24)	-2.085*** (-2.85)	-2.040*** (-2.63)	-2.179*** (-3.02)	-2.078*** (-2.98)	0.884 (1.09)	-1.381* (-1.80)	-1.311* (-1.65)	-2.100*** (-2.96)
LΔMIPI	-0.728 (-1.38)	-0.746 (-1.39)	-0.893 (-1.52)	-0.847 (-1.50)	-0.794 (-1.46)	-0.755* (-1.73)	-1.000* (-1.70)	-1.041* (-1.72)	-0.685 (-1.38)
LΔEXCH	0.235 (1.44)	0.220 (1.36)	0.220 (1.32)	0.206 (1.26)	0.190 (1.21)	0.194 (1.44)	0.188 (1.20)	0.222 (1.36)	0.165 (1.25)
LΔOILP	-0.023 (-0.90)	-0.023 (-0.93)	-0.028 (-1.05)	-0.030 (-1.11)	-0.034 (-1.23)	-0.033 (-1.22)	-0.033 (-1.18)	-0.028 (-1.04)	-0.019 (-0.76)
LEPU	0.027** (1.99)								
LEPUNews		0.022** (2.00)							
LEPUMON			0.007 (1.30)						
LEPUFISC				0.007 (1.03)					
LFRU					0.008** (2.07)				
LMU						0.251*** (4.91)			
LFU							0.053** (2.49)		
LVXO								0.026** (2.31)	
LSP500RV									0.199* (1.82)
<i>N</i>	382	382	382	382	379	382	382	382	382
<i>RMSE</i>	0.039	0.039	0.039	0.039	0.039	0.036	0.039	0.038	0.037
<i>R</i> ²	0.415	0.415	0.404	0.403	0.415	0.481	0.415	0.422	0.452
<i>Adj. R</i> ²	0.402	0.402	0.391	0.390	0.402	0.470	0.403	0.409	0.440

t statistics in parentheses, * p < 0.10, ** p < 0.05, *** p < 0.01. *L* denotes the lagged value for the explanatory variable.

Table 5: Impact of Unobservable Macroeconomic and Financial Uncertainty Indexes on the Volatility of the Commodity Futures Price Index.

	(1)	(2)	(3)	(4)	(5)	(6)
MU1	0.300*** (5.11)					
MU3		0.359*** (5.27)				
MU12			0.661*** (4.63)			
FU1				0.108*** (2.58)		
FU3					0.138** (2.55)	
FU12						0.364** (2.45)
<i>N</i>	384	384	384	384	384	384
<i>RMSE</i>	0.039	0.039	0.040	0.046	0.046	0.047
<i>R</i> ²	0.391	0.399	0.360	0.160	0.156	0.140
<i>Adj. R</i> ²	0.390	0.398	0.359	0.158	0.154	0.137

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 6: Impact of Macroeconomic Uncertainty (MU) and Finance Uncertainty (FU) Indexes on the Realized Volatility of the Commodity price index when controlling for macroeconomic and stock-market factors of commodity market volatility.

	(1)	(2)	(3)	(4)	(5)	(6)
$\Delta SP500$	-0.093* (-1.68)	-0.088 (-1.62)	-0.101* (-1.78)	-0.136* (-1.79)	-0.138* (-1.81)	-0.145* (-1.88)
TERM	0.126 (0.63)	0.172 (0.84)	0.128 (0.57)	0.009 (0.04)	0.015 (0.06)	0.043 (0.17)
$\Delta MEMP$	0.314 (0.33)	0.426 (0.44)	-0.480 (-0.49)	-3.135*** (-3.35)	-3.165*** (-3.38)	-3.301*** (-3.51)
ΔMIP	-1.711*** (-3.10)	-1.644*** (-3.10)	-1.614*** (-3.02)	-2.134*** (-2.78)	-2.138*** (-2.77)	-2.156*** (-2.74)
MU1	0.250*** (5.66)					
MU3		0.304*** (5.61)				
MU12			0.512*** (4.48)			
FU1				0.059*** (3.14)		
FU3					0.074*** (3.08)	
FU12						0.183*** (2.82)
N	383	383	383	383	383	383
$RMSE$	0.038	0.038	0.039	0.041	0.041	0.041
R^2	0.435	0.439	0.407	0.330	0.328	0.321
$Adj. R^2$	0.427	0.431	0.399	0.321	0.319	0.312

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 7: Impact of unobservable Macroeconomic Uncertainty (MU) and Finance Uncertainty (FU) Indexes on the Realized Volatility of the Commodity price index when controlling for macroeconomic factors, exchange rates and oil shocks.

	(1)	(2)	(3)	(4)	(5)	(6)
$\Delta SP500$	-0.087 (-1.55)	-0.081 (-1.47)	-0.092 (-1.62)	-0.126* (-1.82)	-0.128* (-1.84)	-0.135* (-1.92)
TERM	0.122 (0.63)	0.167 (0.86)	0.127 (0.59)	0.008 (0.03)	0.014 (0.06)	0.040 (0.16)
$\Delta MEMP$	0.076 (0.09)	0.188 (0.21)	-0.701 (-0.77)	-3.310*** (-3.34)	-3.340*** (-3.36)	-3.474*** (-3.48)
$\Delta MIPI$	-1.578*** (-3.20)	-1.512*** (-3.17)	-1.472*** (-3.08)	-1.965*** (-3.00)	-1.969*** (-2.99)	-1.984*** (-2.96)
$\Delta EXCH$	0.054 (0.44)	0.061 (0.49)	0.087 (0.66)	0.064 (0.48)	0.064 (0.48)	0.063 (0.46)
$\Delta OILP$	-0.067* (-1.66)	-0.066* (-1.68)	-0.069* (-1.72)	-0.079* (-1.72)	-0.079* (-1.72)	-0.080* (-1.72)
MU1	0.244*** (6.43)					
MU3		0.297*** (6.38)				
MU12			0.501*** (5.05)			
FU1				0.057*** (3.25)		
FU3					0.072*** (3.19)	
FU12						0.178*** (2.92)
N	383	383	383	383	383	383
$RMSE$	0.037	0.037	0.038	0.041	0.041	0.041
R^2	0.449	0.453	0.424	0.350	0.348	0.342
$Adj. R^2$	0.439	0.443	0.413	0.338	0.336	0.329

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 8: Impact of Alternative Economic Uncertainty Measures (Both the Observable and the Latent Measures) on the Volatility of the Commodity Price Index for the 2000-2016 financialization period.

$\Delta SP500$ is the log-difference of the S&P500 stock-market index, the TERM is the term spread, $\Delta MIPI$ is the log-difference of the Manufacturing US Industrial Production Index and $\Delta MEMP$ is the log-difference of the Manufacturing US Employment. EPU stands for the Economic Policy Uncertainty. The variables EPUNEWS, EPUFISC, EPUMON and the Financial Regulation Uncertainty (FRU) are the respective components of the Baker *et al.* (2016) EPU index representing the Uncertainty about macroeconomic news, about fiscal policy, monetary policy, and financial regulation uncertainty respectively. The MU and FU series represent the logarithms of the unobservable macroeconomic and financial uncertainty measures. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index. The data sample covers the period from January 2000 till December 2016.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$\Delta SP500$	-0.201* (-1.68)	-0.202* (-1.66)	-0.206* (-1.69)	-0.213* (-1.72)	-0.207* (-1.76)	-0.115 (-1.59)	-0.175 (-1.53)	-0.138 (-1.32)	0.034 (0.57)
TERM	0.163 (0.51)	0.324 (1.03)	0.434 (1.38)	0.440 (1.37)	0.187 (0.69)	0.575** (2.29)	0.236 (0.80)	0.204 (0.73)	0.077 (0.37)
$\Delta MEMP$	-4.324*** (-4.17)	-4.228*** (-3.91)	-4.134*** (-3.79)	-4.251*** (-3.99)	-4.174*** (-4.02)	1.690 (1.24)	-3.154*** (-3.18)	-2.470** (-2.16)	-2.306*** (-2.78)
$\Delta MIPI$	-2.258** (-2.21)	-2.334** (-2.25)	-2.403** (-2.30)	-2.424** (-2.34)	-2.252** (-2.33)	-1.799*** (-3.00)	-2.308** (-2.36)	-2.362** (-2.44)	-1.327** (-2.12)
EPU	0.021** (2.34)								
EPUNEWS		0.010 (1.43)							
EPUMON			0.003 (0.80)						
EPUFISC				0.000 (0.00)					
FRU					0.007 (1.42)				
MU						0.355*** (4.94)			
FU							0.059** (2.10)		
VXO								0.033** (2.02)	
SP500RV									0.463*** (8.15)
<i>N</i>	204	204	204	204	203	204	204	204	204
<i>RMSE</i>	0.039	0.039	0.039	0.039	0.039	0.032	0.039	0.038	0.031
<i>R</i> ²	0.421	0.412	0.409	0.408	0.417	0.606	0.425	0.442	0.639
<i>Adj. R</i> ²	0.406	0.397	0.394	0.393	0.402	0.596	0.411	0.428	0.629

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 9: Impact of Alternative Economic Uncertainty Measures on the Volatility of the Commodity Price Index for the pre-financialization (1988-1999) period.

$\Delta SP500$ is the log-difference of the S&P500 stock-market index, the TERM is the term spread, $\Delta MIPI$ is the log-difference of the Manufacturing US Industrial Production Index and $\Delta MEMP$ is the log-difference of the Manufacturing US Employment. EPU stands for the Economic Policy Uncertainty. The variables EPUNEWS, EPUFISC, EPUMON and the Financial Regulation Uncertainty (FRU) are the respective components of the Baker *et al.* (2016) EPU index representing the Uncertainty about macroeconomic news, about fiscal policy, monetary policy, and financial regulation uncertainty respectively. The MU and FU series represent the logarithms of the unobservable macroeconomic and financial uncertainty measures. The VXO is the logarithm of the VXO index, and the SP500RV is the monthly realized variance of the daily returns of the S&P 500 index. The data sample covers the period from January 1988 till December 1999.

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)
$\Delta SP500$	-0.049 (-0.66)	-0.036 (-0.46)	-0.028 (-0.34)	-0.053 (-0.69)	-0.048 (-0.62)	-0.056 (-0.70)	-0.049 (-0.61)	-0.025 (-0.31)	-0.037 (-0.38)
TERM	-0.942** (-2.49)	-0.800*** (-2.65)	-0.730*** (-2.68)	-0.918*** (-2.59)	-0.678** (-2.49)	-0.401 (-1.59)	-0.315 (-1.40)	-0.153 (-0.53)	-0.484** (-2.19)
$\Delta MEMP$	-1.303 (-1.00)	-1.113 (-0.88)	-1.975 (-1.25)	-1.838 (-1.20)	-2.243 (-1.26)	-1.108 (-0.79)	-1.759 (-1.04)	-2.222 (-1.25)	-2.599 (-1.30)
$\Delta MIPI$	-0.858 (-1.18)	-1.002 (-1.26)	-0.677 (-1.00)	-0.655 (-1.04)	-0.778 (-1.10)	-1.141 (-1.25)	-1.392 (-1.36)	-1.366 (-1.37)	-1.201 (-1.22)
EPU	0.051 (1.49)								
EPUNEWS		0.051 (1.52)							
EPUMON			0.019 (1.56)						
EPUFISC				0.022 (1.57)					
FRU					0.009 (1.48)				
MU						0.151 (1.51)			
FU							0.060** (1.99)		
VXO								0.037* (1.90)	
SP500RV									0.056 (1.02)
<i>N</i>	179	179	179	179	176	179	179	179	179
<i>RMSE</i>	0.041	0.040	0.041	0.041	0.042	0.042	0.042	0.041	0.042
<i>R</i> ²	0.131	0.151	0.121	0.126	0.115	0.104	0.104	0.133	0.081
<i>Adj. R</i> ²	0.105	0.127	0.095	0.101	0.089	0.078	0.078	0.108	0.055

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 10: Impact of Economic Policy Uncertainty (EPU) Index on the volatility of agricultural, metals and energy commodity futures prices

In this table we report the estimated beta coefficients for the bivariate regression models in which we regress Economic Policy Uncertainty (EPU) on the volatility of agricultural, metals and energy commodity prices.

Panel A: Agricultural Commodity Markets						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
EPU	0.028** (2.01)	0.032 (1.59)	0.047*** (2.73)	0.007 (0.46)	0.083*** (3.05)	0.046** (2.46)
<i>N</i>	384	384	384	384	384	384
<i>RMSE</i>	0.057	0.052	0.044	0.043	0.093	0.060
<i>R</i> ²	0.019	0.029	0.081	0.002	0.059	0.044
<i>Adj. R</i> ²	0.017	0.027	0.079	-0.001	0.056	0.041

Panel B: Energy Commodity Markets				
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum
EPU	0.140** (2.18)	0.097** (1.98)	0.091* (1.68)	0.127** (2.39)
<i>N</i>	360	384	348	384
<i>RMSE</i>	0.146	0.122	0.119	0.127
<i>R</i> ²	0.069	0.048	0.046	0.074
<i>Adj. R</i> ²	0.067	0.045	0.043	0.071

Panel C: Metals Commodity Markets				
	(1) copper	(2) gold	(3) silver	(4) platinum
EPU	0.044 (1.24)	0.028** (2.49)	0.079** (2.21)	0.045*** (2.61)
<i>N</i>	384	384	384	384
<i>RMSE</i>	0.078	0.027	0.087	0.046
<i>R</i> ²	0.024	0.080	0.061	0.070
<i>Adj. R</i> ²	0.022	0.078	0.058	0.067

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 11: Impact of JLN Macroeconomic Uncertainty (MU) Index on the volatility of agricultural, metals and energy commodity futures prices

In this table we report the estimated beta coefficients for the bivariate regression models in which we regress the JLN Macroeconomic Uncertainty (MU) on the volatility of agricultural, metals and energy commodity prices.

Panel A: Agricultural Commodity Markets						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
MU	0.205*** (5.58)	0.241*** (2.94)	0.213*** (4.46)	0.168** (2.57)	0.235*** (3.17)	0.296*** (4.29)
<i>N</i>	384	384	384	384	384	384
<i>RMSE</i>	0.054	0.048	0.042	0.041	0.094	0.056
<i>R</i> ²	0.101	0.162	0.166	0.117	0.047	0.181
<i>Adj. R</i> ²	0.099	0.160	0.163	0.115	0.044	0.179

Panel B: Energy Commodity Markets				
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum
MU	0.732*** (4.17)	0.559*** (6.82)	0.717*** (4.46)	0.647*** (4.99)
<i>N</i>	360	384	348	384
<i>RMSE</i>	0.136	0.115	0.102	0.119
<i>R</i> ²	0.193	0.156	0.298	0.188
<i>Adj. R</i> ²	0.191	0.153	0.296	0.186

Panel C: Metals Commodity Markets				
	(1) copper	(2) gold	(3) silver	(4) platinum
MU	0.385*** (2.71)	0.158*** (3.92)	0.346** (2.35)	0.278*** (4.15)
<i>N</i>	384	384	384	384
<i>RMSE</i>	0.071	0.024	0.085	0.041
<i>R</i> ²	0.187	0.247	0.116	0.265
<i>Adj. R</i> ²	0.184	0.245	0.114	0.263

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 12: Impact of JLN Financial Uncertainty (FU) Index on the volatility of agricultural, metals and energy commodity futures prices

In this table we report the estimated beta coefficients for the bivariate regression models in which we regress Financial Uncertainty (FU) on the monthly realized variance of agricultural, metals and energy commodity futures prices.

Panel A: Agricultural Commodity Markets						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
FU	0.096*** (2.79)	0.087* (1.68)	0.103*** (3.06)	0.057 (1.43)	0.133*** (3.38)	0.105** (2.05)
<i>N</i>	384	384	384	384	384	384
<i>RMSE</i>	0.055	0.051	0.044	0.043	0.094	0.060
<i>R</i> ²	0.059	0.057	0.103	0.035	0.040	0.060
<i>Adj. R</i> ²	0.057	0.054	0.101	0.033	0.037	0.058

Panel B: Energy Commodity Markets				
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum
FU	0.353*** (3.10)	0.246*** (3.85)	0.299*** (2.58)	0.297*** (3.32)
<i>N</i>	360	384	348	384
<i>RMSE</i>	0.142	0.120	0.113	0.125
<i>R</i> ²	0.119	0.080	0.127	0.105
<i>Adj. R</i> ²	0.116	0.078	0.125	0.103

Panel C: Metals Commodity Markets				
	(1) copper	(2) gold	(3) silver	(4) platinum
FU	0.123 (1.33)	0.059** (2.07)	0.121 (1.31)	0.139*** (3.33)
<i>N</i>	384	384	384	384
<i>RMSE</i>	0.077	0.027	0.088	0.043
<i>R</i> ²	0.051	0.092	0.038	0.176
<i>Adj. R</i> ²	0.048	0.090	0.035	0.174

t statistics in parentheses, * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 13: Impact of Economic Policy Uncertainty (EPU) on the volatility of agricultural, metals and energy prices when controlling for macroeconomic fundamentals

Panel A: Agricultural Commodity Markets						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
EPU	0.011 (0.70)	0.016 (1.01)	0.023 (1.42)	-0.015 (-1.31)	0.069** (2.18)	0.024 (1.33)
ΔSP500	0.005 (0.07)	-0.182* (-1.84)	-0.153** (-2.18)	-0.126 (-1.43)	-0.065 (-0.82)	-0.121 (-1.35)
TERM	0.295 (0.68)	0.153 (0.37)	0.811*** (2.62)	0.639* (1.89)	0.424 (0.59)	0.485 (1.08)
ΔMEMP	-4.866*** (-4.52)	-0.827 (-0.50)	-2.114* (-1.80)	-1.323 (-1.11)	-0.517 (-0.24)	-1.098 (-0.69)
ΔMIPI	-0.554 (-1.04)	-1.861** (-2.11)	-0.775 (-1.40)	-1.342** (-2.15)	-0.594 (-0.69)	-2.113*** (-2.76)
<i>N</i>	383	383	383	383	383	383
<i>RMSE</i>	0.054	0.050	0.042	0.041	0.092	0.058
<i>R</i> ²	0.117	0.119	0.179	0.101	0.063	0.120
<i>Adj. R</i> ²	0.105	0.107	0.168	0.090	0.050	0.108
Panel B: Energy Commodity Markets						
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum		
EPU	0.115** (2.20)	0.100* (1.94)	0.053 (1.50)	0.122** (2.40)		
ΔSP500	-0.191 (-1.07)	-0.047 (-0.32)	-0.470*** (-2.66)	-0.102 (-0.65)		
TERM	-1.521* (-1.67)	-1.983** (-2.32)	-0.746 (-1.03)	-1.888** (-2.12)		
ΔMEMP	-11.259*** (-4.26)	-9.378*** (-4.69)	-9.943*** (-4.63)	-9.245*** (-4.17)		
ΔMIPI	-4.164** (-2.11)	-1.716* (-1.77)	-3.586** (-1.97)	-2.950* (-1.90)		
<i>N</i>	360	383	348	383		
<i>RMSE</i>	0.133	0.115	0.105	0.118		
<i>R</i> ²	0.232	0.167	0.254	0.209		
<i>Adj. R</i> ²	0.221	0.156	0.243	0.198		
Panel C: Metals Commodity Markets						
	(1) copper	(2) gold	(3) silver	(4) platinum		
EPU	0.028 (1.45)	0.021*** (2.99)	0.056** (2.04)	0.029*** (2.89)		
ΔSP500	-0.381* (-1.74)	-0.128*** (-2.61)	-0.390* (-1.88)	-0.184** (-2.13)		
TERM	-0.325 (-0.56)	-0.152 (-0.77)	0.386 (0.51)	0.067 (0.20)		
ΔMEMP	-4.867 (-1.42)	-0.514 (-0.89)	1.706 (0.78)	-1.253 (-1.50)		
ΔMIPI	-0.859 (-1.11)	-1.284** (-2.46)	-2.754** (-2.05)	-1.764* (-1.74)		
<i>N</i>	383	383	383	383		
<i>RMSE</i>	0.074	0.025	0.084	0.043		
<i>R</i> ²	0.134	0.236	0.124	0.184		
<i>Adj. R</i> ²	0.122	0.226	0.112	0.173		

t statistics in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 14: Impact of Macroeconomic Uncertainty (MU) on the volatility of agricultural, metals and energy prices when controlling for macroeconomic fundamentals

Panel A: Agricultural Commodity Markets						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
MU	0.121** (2.33)	0.262*** (3.70)	0.218*** (4.18)	0.167** (2.53)	0.327*** (3.46)	0.347*** (4.96)
ΔSP500	0.037 (0.49)	-0.109 (-1.57)	-0.097* (-1.83)	-0.068 (-1.06)	0.002 (0.03)	-0.026 (-0.50)
TERM	0.455 (1.27)	0.419 (1.12)	1.143*** (3.64)	0.529* (1.87)	1.298** (2.11)	0.863** (2.25)
ΔMEMP	-2.898* (-1.93)	3.480*** (2.66)	1.428 (0.93)	1.526 (1.39)	4.659** (2.07)	4.606*** (2.70)
ΔMIPI	-0.463 (-1.10)	-1.602*** (-2.59)	-0.644 (-1.45)	-0.963** (-2.28)	-0.685 (-0.72)	-1.788*** (-3.60)
N	383	383	383	383	383	383
RMSE	0.053	0.047	0.040	0.040	0.091	0.054
R ²	0.132	0.210	0.251	0.152	0.077	0.236
Adj. R ²	0.121	0.200	0.241	0.141	0.065	0.226
Panel B: Energy Commodity Markets						
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum		
MU	0.469*** (3.52)	0.431*** (4.29)	0.565*** (4.16)	0.500*** (4.04)		
ΔSP500	-0.099 (-0.59)	0.038 (0.26)	-0.300** (-2.07)	-0.006 (-0.04)		
TERM	-0.068 (-0.11)	-0.732 (-1.46)	0.093 (0.18)	-0.375 (-0.65)		
ΔMEMP	-3.860 (-1.27)	-2.599 (-1.01)	-0.520 (-0.19)	-1.399 (-0.52)		
ΔMIPI	-4.436** (-2.34)	-1.914 (-1.60)	-3.098** (-2.17)	-3.225** (-2.25)		
N	360	383	348	383		
RMSE	0.132	0.114	0.099	0.117		
R ²	0.236	0.174	0.330	0.214		
Adj. R ²	0.225	0.163	0.320	0.203		
Panel C: Metals Commodity Markets						
	(1) copper	(2) gold	(3) silver	(4) platinum		
MU	0.369*** (3.54)	0.157*** (5.22)	0.462*** (3.50)	0.309*** (4.72)		
ΔSP500	-0.281* (-1.66)	-0.089*** (-2.76)	-0.275* (-1.85)	-0.103** (-1.98)		
TERM	0.114 (0.20)	0.134 (0.78)	1.159* (1.73)	0.489 (1.62)		
ΔMEMP	1.174 (0.43)	2.021*** (2.89)	9.182*** (3.69)	3.783** (2.28)		
ΔMIPI	-0.543 (-0.84)	-1.226*** (-3.40)	-2.530*** (-3.07)	-1.542** (-2.39)		
N	383	383	383	383		
RMSE	0.070	0.023	0.080	0.039		
R ²	0.212	0.325	0.204	0.328		
Adj. R ²	0.202	0.316	0.194	0.319		

t statistics in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 15: Impact of Financial Uncertainty (FU) on the volatility of agricultural, metals and energy prices when controlling for macroeconomic fundamentals.

Agricultural Commodities						
	(1) cocoa	(2) corn	(3) cotton	(4) soybeans	(5) sugar	(6) wheat
FU	0.048 (1.57)	0.056 (1.63)	0.069*** (2.82)	0.027 (1.09)	0.131** (2.52)	0.072* (1.83)
DSP500	0.028 (0.38)	-0.157* (-1.75)	-0.123* (-1.93)	-0.103 (-1.29)	-0.021 (-0.25)	-0.090 (-1.15)
TERM	0.379 (1.02)	0.290 (0.67)	1.019*** (3.17)	0.453 (1.52)	1.094 (1.63)	0.693 (1.53)
DMEMP	-3.940*** (-3.39)	0.235 (0.14)	-0.808 (-0.60)	-0.726 (-0.70)	1.835 (0.79)	0.257 (0.14)
DMIPI	-0.669 (-1.44)	-2.024** (-2.41)	-1.007* (-1.85)	-1.228** (-2.11)	-1.242 (-1.53)	-2.347*** (-3.25)
<i>N</i>	383	383	383	383	383	383
<i>RMSE</i>	0.053	0.049	0.041	0.041	0.092	0.057
<i>R</i> ²	0.126	0.131	0.199	0.100	0.060	0.132
<i>Adj. R</i> ²	0.114	0.119	0.188	0.088	0.048	0.121
Panel B: Energy Commodity Markets						
	(1) crudeoil	(2) heatoil	(3) gasoline	(4) petroleum		
FU	0.213*** (4.18)	0.157*** (3.88)	0.132** (2.53)	0.196*** (4.32)		
DSP500	-0.127 (-0.73)	-0.002 (-0.02)	-0.428*** (-2.59)	-0.044 (-0.29)		
TERM	-0.443 (-0.83)	-0.990** (-2.18)	-0.214 (-0.39)	-0.685 (-1.36)		
DMEMP	-7.323*** (-2.85)	-6.627*** (-3.38)	-7.199*** (-3.55)	-5.790*** (-2.74)		
DMIPI	-5.276** (-2.55)	-2.639** (-2.24)	-4.073** (-2.27)	-4.074** (-2.55)		
<i>N</i>	360	383	348	383		
<i>RMSE</i>	0.133	0.116	0.105	0.119		
<i>R</i> ²	0.229	0.152	0.260	0.192		
<i>Adj. R</i> ²	0.218	0.141	0.249	0.181		
Panel C: Metals Commodity Markets						
	(1) copper	(2) gold	(3) silver	(4) platinum		
FU	0.047 (0.80)	0.037** (1.99)	0.091 (1.34)	0.117*** (3.67)		
DSP500	-0.367* (-1.72)	-0.116** (-2.54)	-0.363* (-1.88)	-0.129* (-1.79)		
TERM	-0.045 (-0.07)	0.054 (0.26)	0.937 (1.28)	0.301 (0.90)		
DMEMP	-4.038 (-1.17)	0.147 (0.21)	3.306 (1.33)	0.985 (0.85)		
DMIPI	-1.122 (-1.47)	-1.480*** (-2.98)	-3.270** (-2.48)	-2.064** (-2.28)		
<i>N</i>	383	383	383	383		
<i>RMSE</i>	0.074	0.025	0.085	0.041		
<i>R</i> ²	0.131	0.229	0.116	0.257		
<i>Adj. R</i> ²	0.120	0.219	0.105	0.247		

t statistics in parentheses. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

Table 16: Granger Causality Tests Between Macroeconomic Uncertainty (MU) and Realized Variance in Individual Commodity Prices

This table shows the results of the Granger causality tests between the various economic and financial uncertainty measures and commodity market uncertainty. The tests refer to the multivariate 6-factor VAR mode with 4 lags with the following VAR ordering :[MIPI MEMP MU TERM SP500 RV]. The RV is the Realized Variance of each commodity market. The null hypothesis is that the Independent variable does not Granger cause the Dependent variable. With * , ** and *** we reject the null hypothesis of no causality at the 10%, 5% and 1% level respectively. Panels A, B and C show the estimated results for our Granger causality tests when we use different proxies for agricultural, energy and metals commodity markets respectively. The time series sample covers the period from January 1985 until December 2016.

Panel A: Agricultural Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Cocoa RV	MU	3.38	0.497
Corn RV	MU	18.57***	0.001
Cotton RV	MU	10.88**	0.028
Soybeans RV	MU	6.87	0.143
Sugar RV	MU	17.58***	0.001
Wheat RV	MU	13.72***	0.008
MU	Cocoa RV	3.77	0.438
MU	Corn RV	7.27	0.122
MU	Cotton RV	2.08	0.721
MU	Soybeans RV	4.80	0.308
MU	Sugar RV	6.94	0.139
MU	Wheat RV	21.49***	0.000

Panel B: Energy Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Crude oil RV	MU	14.43***	0.006
Heating oil RV	MU	19.17***	0.001
Petroleum RV	MU	26.67***	0.000
Unleaded gasoline RV	MU	27.67***	0.000
MU	Crude oil RV	14.56***	0.006
MU	Heating oil RV	11.23**	0.024
MU	Petroleum RV	12.07**	0.017
MU	Unleaded gasoline RV	19.59***	0.001

Panel C: Metals Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Copper RV	MU	18.23***	0.001
Gold RV	MU	24.12***	0.000
Platinum RV	MU	37.23***	0.000
Silver RV	MU	18.44***	0.001
MU	Copper RV	18.24***	0.001
MU	Gold RV	7.07	0.132
MU	Platinum RV	9.23*	0.056
MU	Silver RV	4.13	0.389

Table 17: Granger Causality Tests Between Financial Uncertainty (FU) and Realized Variance in Individual Commodity Prices

This table shows the results of the Granger causality tests between FU and the Realized Variance (RV) in agricultural, metals and energy markets. The tests refer to the multivariate 6-factor VAR mode with 4 lags with the following VAR ordering :[MIPI MEMP MU TERM SP500 RV]. The RV is the Realized Variance of each commodity market. The null hypothesis is that the Independent variable does not Granger cause the Dependent variable. With *, ** and *** we reject the null hypothesis of no causality at the 10%, 5% and 1% level respectively. Panels A, B and C show the estimated results for our Granger causality tests when we use different proxies for agricultural, energy and metals commodity markets respectively. The time series sample covers the period from January 1985 until December 2016.

Panel A: Agricultural Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Cocoa RV	FU	9.67**	0.046
Corn RV	FU	8.28*	0.082
Cotton RV	FU	19.14***	0.001
Soybeans RV	FU	1.95	0.744
Sugar RV	FU	10.80**	0.029
Wheat RV	FU	7.27	0.122
FU	Cocoa RV	1.881	0.770
FU	Corn RV	4.71	0.318
FU	Cotton RV	3.01	0.557
FU	Soybeans RV	5.29	0.258
FU	Sugar RV	11.90**	0.018
FU	Wheat RV	9.96**	0.041

Panel B: Energy Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Crude oil RV	FU	5.03	0.284
Heating oil RV	FU	4.87	0.301
Petroleum RV	FU	7.35	0.118
Unleaded gasoline RV	FU	7.26	0.123
FU	Crude oil RV	4.66	0.324
FU	Heating oil RV	4.44	0.349
FU	Petroleum RV	3.41	0.491
FU	Unleaded gasoline RV	3.43	0.488

Panel C: Metals Commodity Markets

Dependent variable	Independent variable	chi-square	p-value
Copper RV	FU	12.73**	0.013
Gold RV	FU	20.67***	0.000
Platinum RV	FU	29.03***	0.000
Silver RV	FU	20.16***	0.000
FU	Copper RV	13.42***	0.009
FU	Gold RV	6.43	0.170
FU	Platinum RV	6.64	0.156
FU	Silver RV	4.64	0.326

Figure 1. Orthogonalized Impulse Response Functions (OIRFs) between Realized Variance in the Commodity Market Index and Macroeconomic Uncertainty (MU)

This figure shows the estimated Orthogonalized Impulse Response Functions between the logarithm of Macroeconomic Uncertainty (MU) and the monthly Realized Variance of the Commodity Futures Market Index (COMRV). The OIRFs are estimated for the baseline 6-factor VAR model which is described in Equations (2) and (3) of the paper. The estimated OIRFs are presented in percentages. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

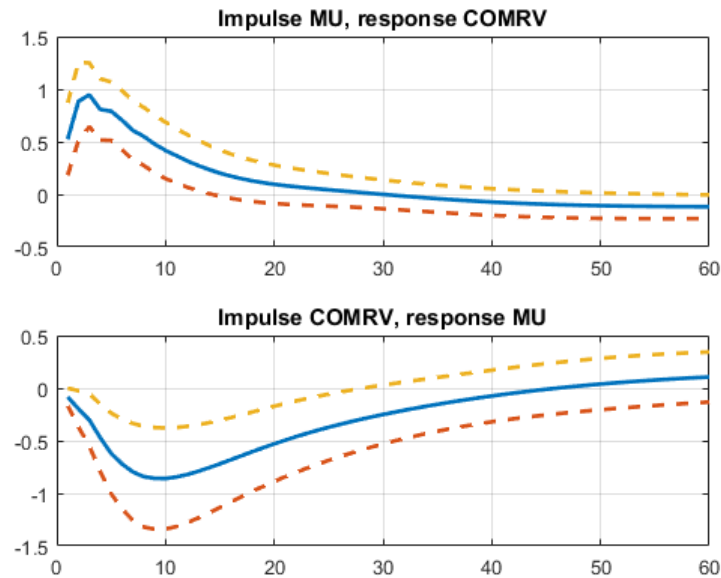


Figure 2. Orthogonalized Impulse Response Functions (OIRFs) between Realized Variance in the Commodity Market Index and Financial Uncertainty (FU)

This figure shows the estimated Orthogonalized Impulse Response Functions between the logarithm of Financial Uncertainty (FU) and the volatility of the Commodity Futures Market Index (COMRV). The OIRFs are estimated for the baseline 6-factor VAR model which is described in Equations (2) and (3) of the paper. The estimated OIRFs are presented in percentages. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

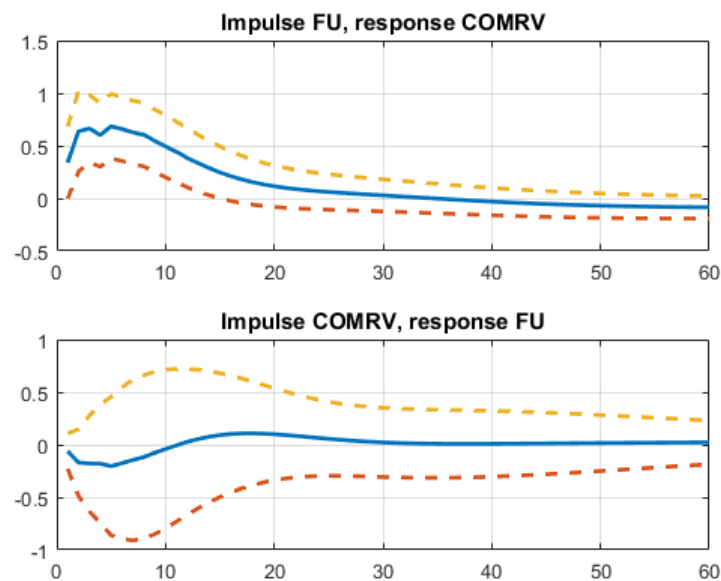


Figure 3. Orthogonalized Impulse Response Functions (OIRFs) between Realized Variance in the Commodity Market Index and Macroeconomic Uncertainty (MU) (alternative VAR ordering)

This figure shows the estimated Orthogonalized Impulse Response Functions between the logarithm of Macroeconomic Uncertainty (MU) and the volatility of the Commodity Futures Market Index (COMRV). The OIRFs are estimated for the 6-factor VAR model with the alternative VAR ordering which is described in Equation (1) of the Appendix. The estimated OIRFs are presented in percentages. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

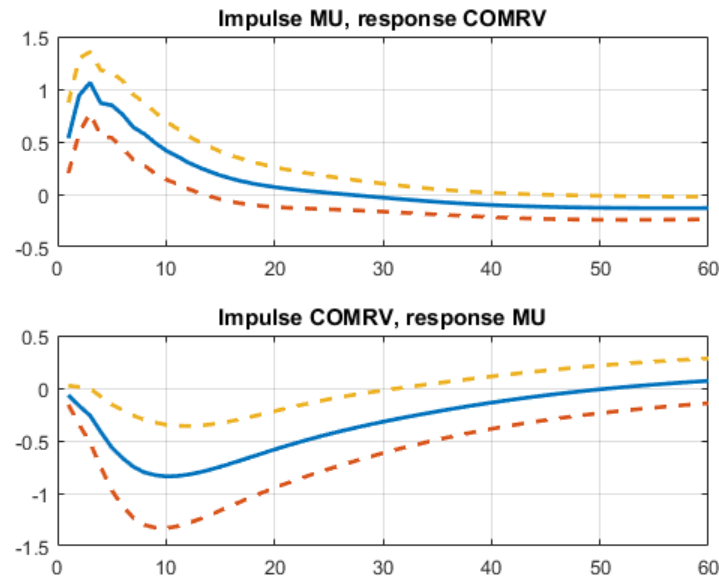


Figure 4. Orthogonalized Impulse Response Functions (OIRFs) between Realized Variance in the Commodity Market Index and Financial Uncertainty (FU) (alternative VAR ordering)

This figure shows the estimated Orthogonalized Impulse Response Functions between the logarithm of Financial Uncertainty (FU) and the volatility of the Commodity Futures Market Index (COMRV). The OIRFs are estimated for the 6-factor VAR model with the alternative VAR ordering which is described in Equation (1) of the Appendix. The estimated OIRFs are presented in percentages. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

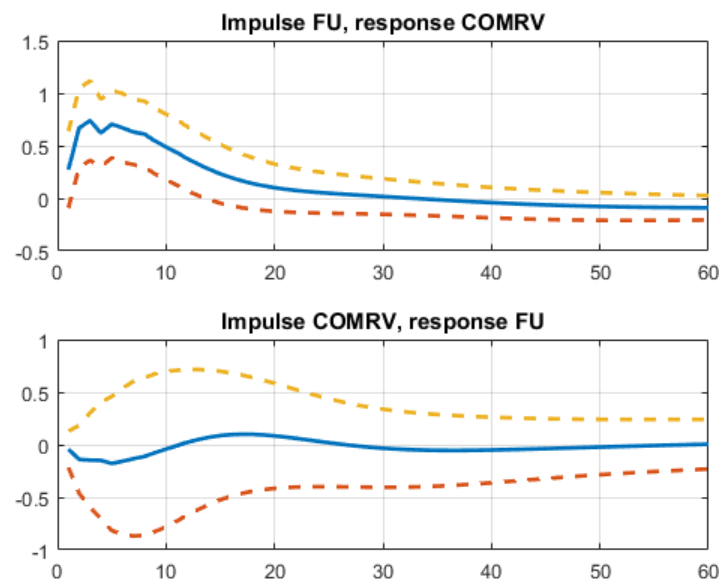


Figure 5. Impulse Response Functions (IRFs) between Realized Variance in Commodity Market Index and Macroeconomic Uncertainty (MU) (8-factor VAR Model)

This figure shows the estimated Impulse Response Functions between the logarithm of Macroeconomic Uncertainty (MU) and the monthly Realized Variance of the Commodity Futures Market Index (COMRV). The IRFs are estimated for the 8-factor VAR model given in Equation (2) of the Appendix. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

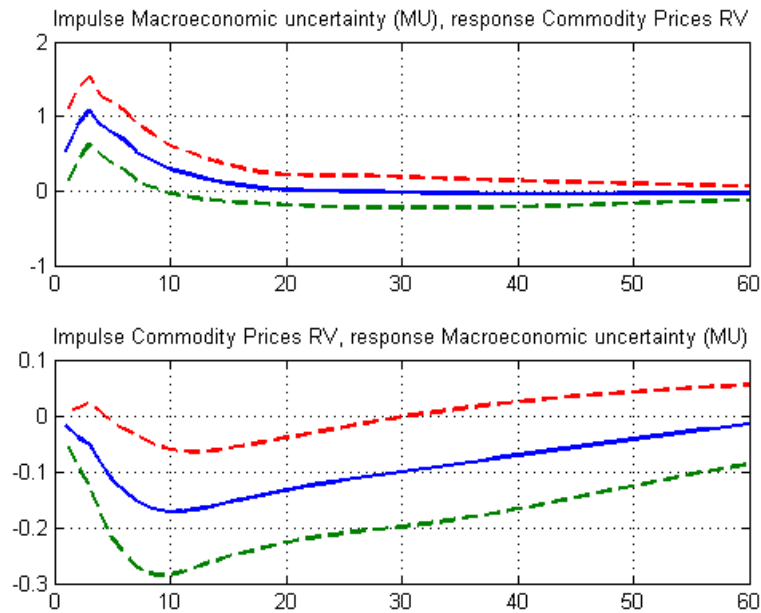


Figure 6. Impulse Response Functions (IRFs) between Realized Variance in Commodity Market Index and Financial Uncertainty (FU) (8-factor VAR Model)

This figure shows the estimated Impulse Response Functions between the Financial Uncertainty (FU) and the monthly Realized Variance of the Commodity Futures Market Index (COMRV). The IRFs are estimated for the 8-factor VAR model given in Equation (2) of the Appendix. The blue line shows the estimated IRFs and the dashed lines show the corresponding 95% bootstrapped confidence intervals. The bootstrapped standard errors have been estimated by using 1000 replications.

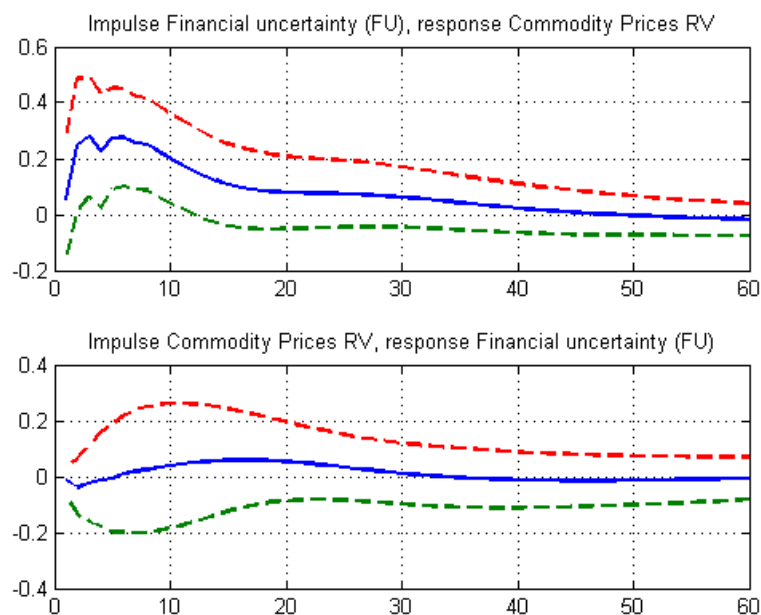


Figure 7. Impulse Response Functions (IRFs) between Realized Variance in Agricultural Commodity Markets and Financial Uncertainty (FU)

This figure shows the estimated IRFs between various measures of economic policy uncertainty and the Realized Variance (RV) in agricultural commodity markets. Panel A shows the responses of Commodity Market Volatility shocks to Macroeconomic Uncertainty Shocks and Panel B shows the responses of Macroeconomic Uncertainty to Agricultural Commodity Volatility Shocks.

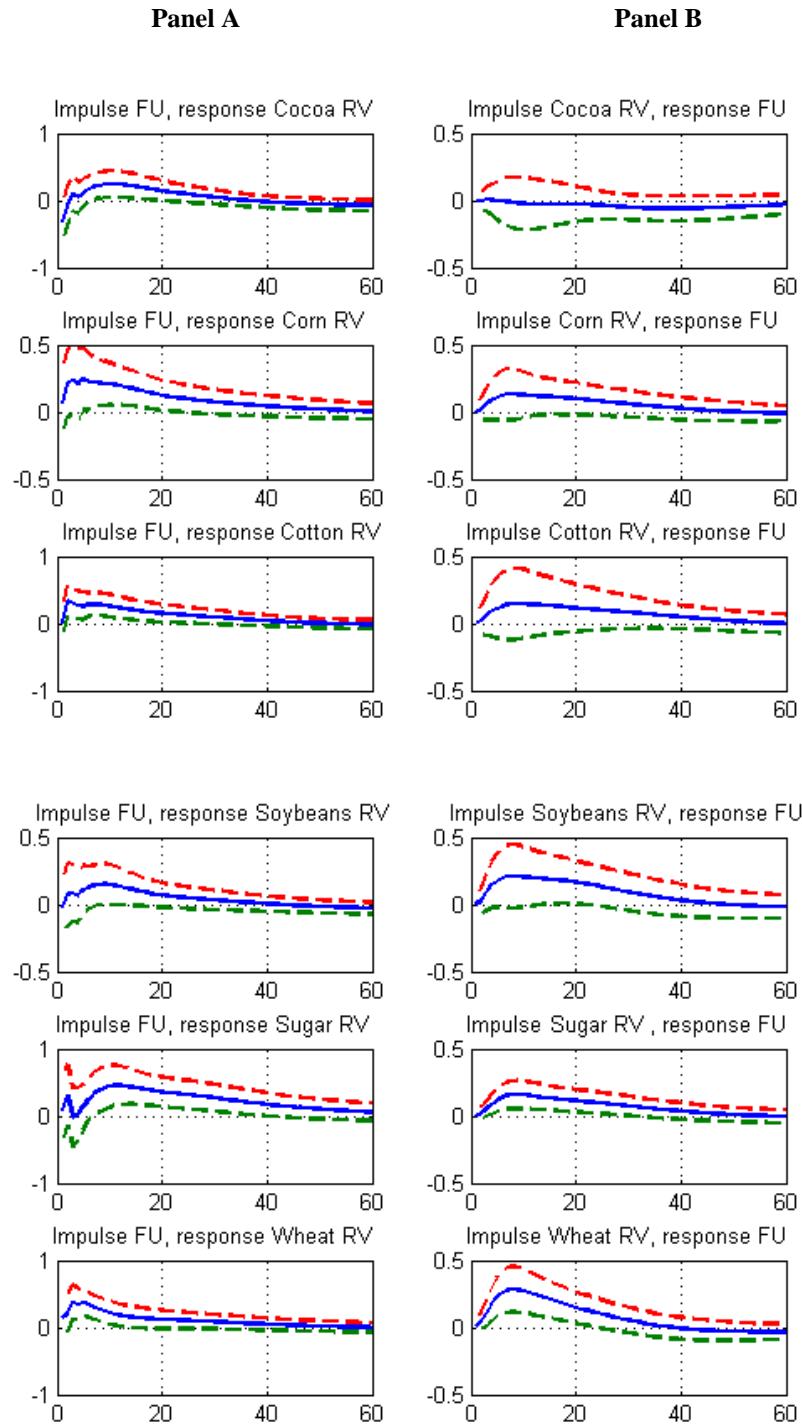


Figure 8. Impulse Response Functions (IRFs) between Realized Variance in Energy Commodity Markets and Financial Uncertainty (FU)

This figure shows the estimated IRFs between various measures of economic policy uncertainty and the Realized Variance (RV) in the energy commodity futures. Panel A shows the responses of Commodity Market Volatility shocks to Macroeconomic Uncertainty Shocks and Panel B shows the responses of Macroeconomic Uncertainty (MU) to energy commodity Volatility Shocks.

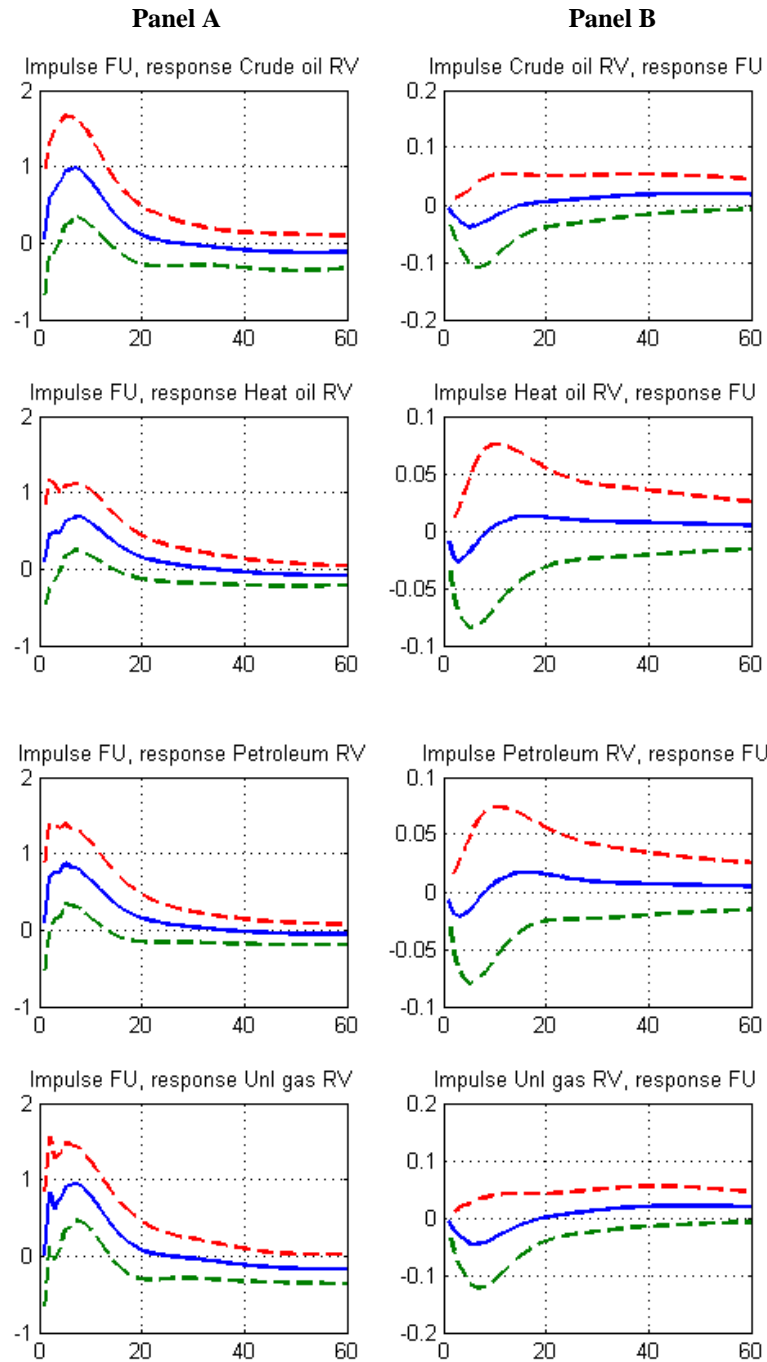
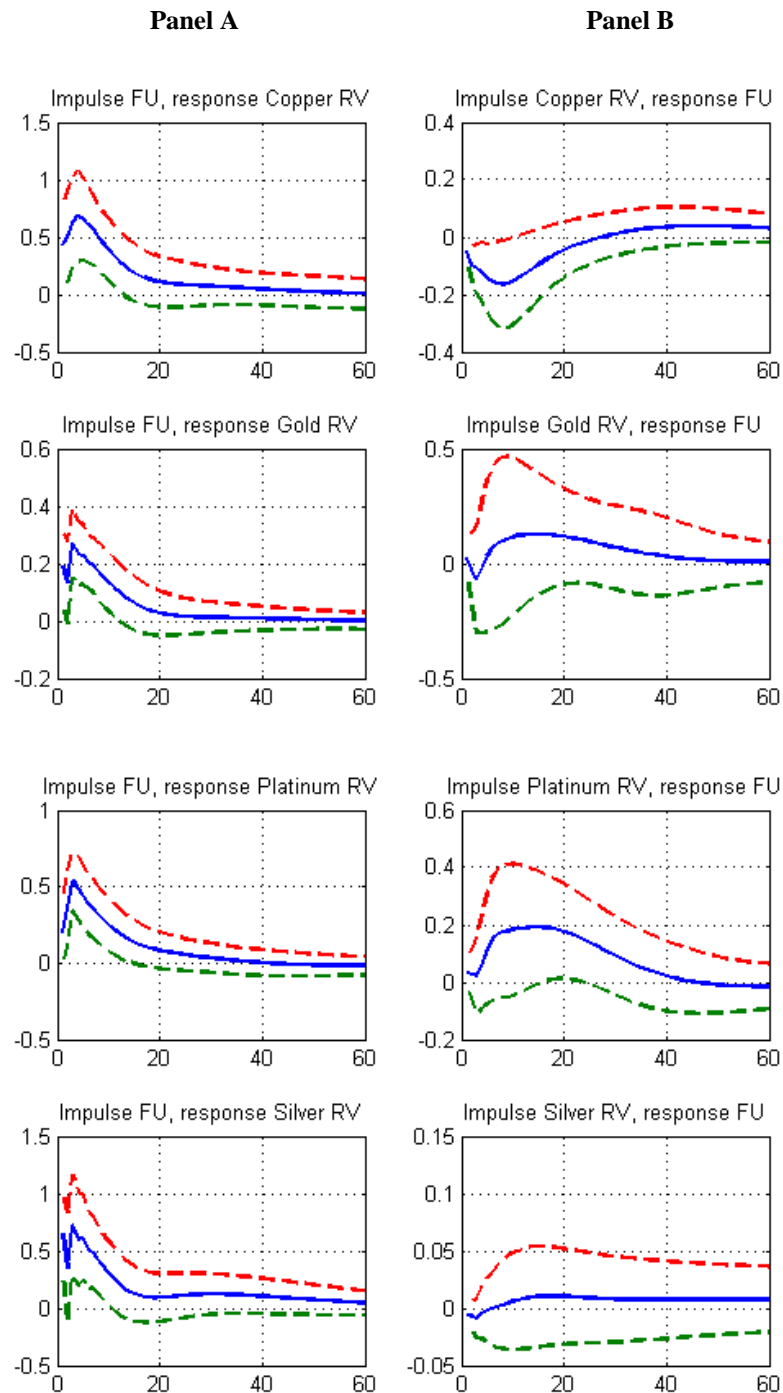


Figure 9. Impulse Response Functions (IRFs) between Realized Variance in Mineral (Metals) Commodity Markets and Financial Uncertainty (FU)

This figure shows the estimated IRFs between various measures of economic policy uncertainty and the Realized Variance (RV) in the mineral markets. Panel A shows the responses of Commodity Market Volatility shocks to Macroeconomic Uncertainty Shocks and Panel B shows the responses of Macroeconomic Uncertainty to Metals Volatility Shocks.



DISCUSSION PAPERS IN ECONOMICS

- 2018/1 Dimitrios Bakas and Athanasios Triantafyllou, *The Impact of Uncertainty Shocks on the Volatility of Commodity Prices.*
- 2017/9 Eghosa Igudia, Rob Ackrill and Simeon Coleman, *Entrepreneurial Responses to Austerity: The Role of the Informal Sector.*
- 2017/8 Rahmi Cetin and Robert Ackrill, *Openness and Growth in Challenging Times: Analysing the Trade-Growth Nexus in Slovakia.*
- 2017/7 Dimitrios Bakas and Yousef Makhoulouf, *Can the Insider-Outsider Theory explain Unemployment Hysteresis in the OECD Countries?*
- 2017/6 Dimitrios Bakas, Georgios Chortareas and Georgios Magkonis, *Volatility and Growth: A not so Straightforward Relationship.*
- 2017/5 Will Rossiter, *Prospects and Challenges for City Region Devolution in Nottingham and the East Midlands.*
- 2017/4 Chunping Liu and Zhirong Ou, *What determines China's housing price dynamics? New evidence from a DSGE-VAR.*
- 2017/3 Morakinyo O. Adetutu and Thomas G. Weyman-Jones, *Fuel subsidies versus market power: is there a countervailing second best welfare optimum?*
- 2017/2 Robert Mullings, *Do Institutions moderate Globalization's effect on Growth?*
- 2017/1 Anthony J Glass, Karligash Kenjegalieva, Victor Ajayi, Morakinyou Adetutu, Robin C. Sickles, *Relative Winners and Losers from Efficiency Spillovers in Africa with Policy Implications for Regional Integration.*
- 2016/4 Sara Ornati, *International Environmental Agreements and Domestic Politics.*
- 2016/3 Yanhui Zhu, Jingwen Fan and Jon Tucker, *The impact of UK monetary policy on gold price dynamics.*
- 2016/2 John Ebireri and Alberto Paloni, *Bank Development and a Lower Degree of Sophistication and Diversification of Developing Countries' Exports.*
- 2016/1 Robert Mullings and Aruneema Mahabir, *Growth by Destination: The Role of Trade in Africa's Recent Growth Episode.*
- 2015/1 Andrew Atherton, João R. Faria, Dongxu Wu and Zhongmin Wu, *Human Capital, Entrepreneurial Entry and Survival.*
- 2014/5 Leighton Vaughan Williams, *The US Supreme Court and the 'Affordable Care Act': An Exercise in Closed-Door Forecasting.*
- 2014/4 Juan Carlos Cuestas and Barry Harrison, *Unemployment Persistence in the EU15.*
- 2014/3 Francesco Marchionne and Sunny Parekh, *Growth, Debt, and Inequality.*

- 2014/2 Bo Jiang, Zhongmin Wu, Bruce Philp and Simeon Coleman, *Macro Stress Testing in the Banking System of China*.
- 2014/1 Eghosa Igudia, Rob Ackrill, Simeon Coleman and Carlyn Dobson, *Austerity Measures or Entrepreneurial Development? The case of the Nigerian Informal Economy*.
- 2013/5 Barry Harrison and Theodorus Wisnu Widjaja, *Did the Financial Crisis impact on the Capital Structure of Firms?*
- 2013/4 Geetha Ravishankar and Marie Stack, *The Gravity Model and Trade Efficiency: A Stochastic Frontier Analysis of Potential Trade*.
- 2013/3 Chunping Liu and Patrick Minford, *How Important is the Credit Channel? An Empirical Study of the US Banking Crisis*.
- 2013/2 Chunping Liu and Patrick Minford, *Comparing Behavioural and Rational Expectations for the US Post-War Economy*.
- 2013/1 Dan Wheatley, *Is it good to share? Debating patterns in availability and use of job share*.
- 2012/3 Simeon Coleman and Vitor Leone, *Time-Series Characteristics Of UK Commercial Property Returns: Testing For Multiple Changes In Persistence*.
- 2012/2 Otavio Ribeiro de Medeiros and Vitor Leone, *Multiple Changes in Persistence vs. Explosive Behaviour: The Dotcom Bubble*.
- 2012/1 Rob Ackrill and Simeon Coleman, *Inflation Dynamics In Central And Eastern European Countries*.
- 2011/4 Simeon Coleman, *Inflation Dynamics and Poverty Rates: Regional and Sectoral Evidence for Ghana*.
- 2011/3 Dan Wheatley and Zhongmin Wu, *Work, Inequality, And The Dual Career Household*.
- 2011/2 Simeon Coleman and Kavita Sirichand, *Fractional Integration and the Volatility Of UK Interest Rates*.
- 2011/1 Simeon Coleman, *Investigating Business Cycle Synchronization In West Africa*.
- 2010/11 Marie Stack and Eric Pentecost, *A Gravity Model Approach To Estimating Prospective Trade Gains in The EU Accession And Associated Countries*.
- 2010/10 Vitor Leone And Bruce Philp, *Surplus-Value And Aggregate Concentration In The UK Economy, 1987-2009*.
- 2010/9 Robert Ackrill and Adrian Kay, *WTO Regulations and Bioenergy Sustainability Certification – Synergies and Possible Conflicts*.

- 2010/8 Paul Alagidede, Simeon Coleman and Juan Carlos Cuestas, *Persistence Of Inflationary Shocks: Implications For West African Monetary Union Membership*.
- 2010/6 Bruce Philp and Dan Wheatley, *The time scarcity and the dual career household: competing perspectives*
- 2010/5 Juan Carlos Cuestas, Sebastián Freille and Patricio O’Gorman, *The media and public agendas: testing for media effects in Argentina Turing the Kirchner administration*
- 2010/4 Vitor Leone, *From property companies to real estate investment trusts: the impact of economic and property factors in the UK commercial property returns*
- 2010/3 Juan Carlos Cuestas and Paulo José Regis, *Purchasing power parity in OECD countries: nonlinear unit root tests revisited*
- 2010/2 Juan Carlos Cuestas and Bruce Philp, *Exploitation and the class struggle*
- 2010/1 Barry Harrison and Winston Moore, *Nonlinearities in Stock Returns for Some Recent Entrants to the EU*
- 2009/7 Joao R. Faria, Le Wang and Zhongmin Wu, *Debts on debts*
- 2009/6 Juan Carlos Cuestas and Luis A. Gil-Alana, *Unemployment hysteresis, structural changes, non-linearities and fractional integration in Central and Eastern Europe*
- 2009/5 Juan Carlos Cuestas and Javier Ordóñez, *Unemployment and common smooth transition trends in Central and Eastern European Countries*
- 2009/4 Stephen Dobson and Carlyn Ramlogan, *Is there a trade-off between income inequality and corruption? Evidence from Latin America*
- 2009/3 Juan Carlos Cuestas and Luís Alberiko Gil-Alana, *Further evidence on the PPP analysis of the Australian dollar: non-linearities, structural changes and fractional integration*
- 2009/2 Estefanía Mourelle and Juan Carlos Cuestas, *Inflation persistence and asymmetries: Evidence for African countries*
- 2009/1 Juan Carlos Cuestas and Barry Harrison, *Further evidence on the real interest rate parity hypothesis in Central and Eastern European Countries: unit roots and nonlinearities*
- 2008/16 Simeon Coleman, *Inflation persistence in the Franc Zone: evidence from disaggregated prices*
- 2008/15 Juan Carlos Cuestas and Paulo Regis, *Nonlinearities and the order of*

integration of order prices

- 2008/14 Peter Dawson and Stephen Dobson, *The influence of social pressure and nationality on individual decisions: evidence from the behaviour of referees*
- 2008/13 Juan Carlos Cuestas and Barry Harrison, *Testing for stationarity of inflation in Central and Eastern European Countries*
- 2008/12 Juan Carlos Cuestas and Dean Garratt, *Is real GDP per capita a stationary process? Smooth transitions, nonlinear trends and unit root testing*
- 2008/11 Antonio Rodriguez Andres and Carlyn Ramlogan-Dobson, *Corruption, privatisation and the distribution of income in Latin America*
- 2008/10 Stephen Dobson and Carlyn Ramlogan, *Is there an openness Kuznets curve? Evidence from Latin America*
- 2008/9 Stephen Dobson, John Goddard and Frank Stähler, *Effort levels in contests: an empirical application of the Tullock model*
- 2008/8 Juan Carlos Cuestas and Estefania Mourelle, *Nonlinearities in real exchange rate determination: Do African exchange rates follow a random walk?*
- 2008/7 Stephen Dobson and John Goddard, *Strategic behaviour and risk taking in football*
- 2008/6 Joao Ricardo Faria, Juan Carlos Cuestas and Estefania Mourelle, *Entrepreneurship and unemployment: A nonlinear bidirectional causality?*
- 2008/5 Dan Wheatley, Irene Hardill and Bruce Philp, *"Managing" reductions in working hours: A study of work-time and leisure preferences in the UK industry*
- 2008/4 Adrian Kay and Robert Ackrill, *Institutional change in the international governance of agriculture: a revised account*
- 2008/3 Juan Carlos Cuestas and Paulo José Regis, *Testing for PPP in Australia: Evidence from unit root test against nonlinear trend stationarity alternatives*
- 2008/2 João Ricardo Faria, Juan Carlos Cuestas and Luis Gil-Alana, *Unemployment and entrepreneurship: A Cyclical Relation*
- 2008/1 Zhongmin Wu, Mark Baimbridge and Yu Zhu, *Multiple Job Holding in the United Kingdom: Evidence from the British Household Panel Survey*

DISCUSSION PAPERS IN POLITICAL ECONOMY

- 2006/3 Ioana Negru, *On Homogeneity and Pluralism within Economics Schools of Thought*
- 2006/2 David Harvie and Bruce Philp, *Learning and Assessment in a Reading Group Format or Reading Capital... For Marks*
- 2006/1 David Harvie, Bruce Philp and Gary Slater, *Regional Well-Being and 'Social Productivity' in Great Britain'*
- 2004/2 Massimo De Angelis and David Harvie, *Globalisation? No Question: Foreign Direct Investment and Labour Commanded*
- 2004/1 David Harvie, *Value-Production and Struggle in the Classroom, or, Educators Within, Against and Beyond Capital*

DISCUSSION PAPERS IN APPLIED ECONOMICS AND POLICY

- 2007/2 Juan Carlos Cuestas, *Purchasing Power Parity in Central and Eastern European Countries: An Analysis of Unit Roots and Non-linearities*
- 2007/1 Juan Carlos Cuestas and Javier Ordóñez, *Testing for Price Convergence among Mercosur Countries*
- 2006/2 Rahmi Cetin and Robert Ackrill, *Foreign Investment and the Export of Foreign and Local Firms: An Analysis of Turkish Manufacturing*
- 2006/1 Robert Ackrill and Adrian Kay, *The EU Financial Perspective 2007-2013 and the Forces that Shaped the Final Agreement*
- 2004/5 Michael A. Smith, David Paton and Leighton Vaughan-Williams, *Costs, Biases and Betting markets: New evidence*
- 2004/4 Chris Forde and Gary Slater, *Agency Working in Britain: Character, Consequences and Regulation*
- 2004/3 Barry Harrison and David Paton, *Do 'Fat Tails' Matter in GARCH Estimation? Stock market efficiency in Romania and the Czech Republic*
- 2004/2 Dean Garratt and Rebecca Taylor, *Issue-based Teaching in Economics*
- 2004/1 Michael McCann, *Motives for Acquisitions in the UK*
- 2003/6 Chris Forde and Gary Slater, *The Nature and Experience of Agency Working in Britain*
- 2003/5 Eugen Mihaita, *Generating Hypothetical Rates of Return for the Romanian Fully Funded Pension Funds*
- 2003/4 Eugen Mihaita, *The Romanian Pension Reform*

2003/3 Joshy Easaw and Dean Garratt, *Impact of the UK General Election on Total Government Expenditure Cycles: Theory and Evidence*

2003/2 Dean Garratt, *Rates of Return to Owner-Occupation in the UK Housing Market*

2003/1 Barry Harrison and David Paton, *The Evolution of Stock Market Efficiency in a Transition Economy: Evidence from Romania.*